

STRUCTURAL VERSUS BEHAVIORAL  
REMEDIES IN THE DEREGULATION  
OF ELECTRICITY MARKETS:  
AN EXPERIMENTAL INVESTIGATION GUIDED  
BY THEORY AND POLICY CONCERNS

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# Structural versus Behavioral Remedies in the Deregulation of Electricity Markets: An Experimental Investigation Guided by Theory and Policy Concerns\*

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## Abstract

We try to better understand the comparative advantages of structural and behavioral remedies of deregulation in electricity markets, an eminent policy issue for which the experimental evidence is scant and problematic. Specifically, we investigate theoretically and experimentally the effects on competition of introducing a forward market — considered a behavioral remedy by the European Commission. We compare this scenario with the best alternative, the structural remedy of reducing concentration by adding one more competitor by divestiture. Our study contributes to the literature by introducing more realistic cost configurations, by teasing apart competition effect and asset effect, and by investigating competitor numbers that reflect the market concentration in the European electricity industries. Our experimental data suggest that introducing a forward market has a positive effect on the aggregate supply in markets with two or three major competitors, configurations typical for the newly accessed and the old European Union member states, respectively. Introducing a forward market also increases efficiency. In contrast to previous findings, our data furthermore suggest that the effect of introducing a forward market is stronger than adding one more competitor both in markets with two, and particularly three, producers. Our data thus provides evidence that behavioral remedies may be more effective than structural remedies. Our data suggest that competition authorities are well advised, in line with EU law (European Commission, 2006a, p.11), to focus on introducing, and facilitating the proper functioning of, forward markets rather than on lowering market concentration by divestiture.

**Keywords:** economics experiments, market power, competition, forward markets, EU electricity market.

**JEL classification:** C91, D61, L13, L43, L94, Q48.

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## **Abstrakt**

Pokoušíme se lépe porozumět komparativním výhodám strukturálních a behaviorálních opatření deregulace na trzích s elektřinou, což je významná otázka při tvorbě veřejné politiky, pro niž existují limitovaná a problematická data. Teoreticky a experimentálně zkoumáme efekty, jež na konkurenci má zavedení termínového trhu – považovaného Evropskou komisí za behaviorální nápravu. Srovnáváme tento scénář s nejlepší alternativou, strukturálním opatřením, které snižuje koncentraci tím, že přidává dalšího konkurenta pomocí divestitury. Naše studie přispívá do existující literatury zavedením realističtějších konfigurací cen pomocí oddělení efektu konkurence od efektu jmění a pomocí prozkoumání počtů konkurentů, které odráží tržní koncentraci na evropských trzích s elektřinou. Naše experimentální data naznačují, že zavedení termínového trhu má pozitivní dopad na agregovanou nabídku na trzích s dvěma nebo třemi hlavními konkurenty, což jsou konfigurace typické pro nově přistoupivší, respektive staré členy Evropské unie. Zavedení termínového trhu také zvyšuje efektivitu. V protikladu k dřívějším zjištěním naše data dále naznačují, že efekt zavedení termínového trhu je silnější než přidání dalšího konkurenta na trzích s dvěma, a zejména třemi, producenty. Naše data tedy přinášejí důkazy, že behaviorální opatření mohou být efektivnější než opatření strukturální. Naše data naznačují, že úřady na ochranu soutěže by se v souladu s právem EU (Evropská komise, 2006a, str. 11) měly zaměřit na zavedení a podporu správného fungování termínových trhů místo toho, aby se zaměřovaly na snižování tržní koncentrace pomocí divestitur.

## 1. Introduction

Concentration in generator markets remains a major problem in the EU electricity markets. The European Commission (2007a, p.7), for example, concludes: “At the wholesale level, gas and electricity markets remain national in scope, and generally maintain the high level of concentration of the pre-liberalization period. This gives scope for exercising market power.” The European Commission suggests structural remedies<sup>2</sup> such as divestiture or asset swaps of power plants on a European scale (2007a, p.15), blocking mergers (2007a, p.12), auctioning large scale Virtual Power Plants (2007a, p.12), stimulating the entrance of new electricity generators (2007a, p.16), and increasing competition by enabling generators from abroad to sell electricity over cross-border transmission lines (2007a, p.8).

Several EU member states have experience with some of these structural remedies. For example, at the end of the nineties, the UK forced dominant electricity generators to divest plants; the two dominant electricity generators NationalPower and PowerGen together divested 6GW in 1996 and another 8GW in 1999, thus lowering concentration (Green, 2006). However, beginning in 2000, the UK experienced mergers which reversed that trend.<sup>3</sup> The UK also experienced a considerable degree of new entry.<sup>4</sup> Belgium, France, Italy, Denmark, and the Netherlands are using, or used in the past, the auctioning of Virtual Power Plants<sup>5</sup> to lower market power (Willems, 2006). Finally, several countries increased the capacity of cross-border transmission lines and harmonized their market regimes with neighboring countries to make it easier for generators to sell electricity over borders, thus increasing competition.

The encouragement of cross-border trading – while creating a larger, European, market – is likely to alleviate the concentration problem only marginally; many electricity companies have merged across borders, and have thus become players in neighboring countries (Matthes, Grashof, and Gores, 2007). Increasing competition is therefore done most efficiently - avoiding duplication of investment in generation assets<sup>6</sup> - by divestiture; enforcing big incumbent power companies to sell parts of their plants, and thus adding to the capacity of competing new entrants. Of interest are

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<sup>2</sup> The European Commission (2006b, p.6) defines structural remedies as “changes to the structure of an undertaking. The most obvious one is the divestiture of an existing business.”

<sup>3</sup> In 2002 one of the largest generators, PowerGen, merged with TXU Europe, thus adding 3GW to its capacity (Green, 2006).

<sup>4</sup> The policy of allowing distributors to sign long-term contracts with independent power producers promoted entry of new electricity producers, mainly with new Combined-Cycle-Gas-Turbine (CCGT) generation technology (Newbery, 2002).

<sup>5</sup> When a generator sells a Virtual Power Plant, he sells part of his production capacity to other generators. This divestiture of generation capacity is called virtual as no production capacity changes hands, and the selling generator remains the owner of all its generation plants (Willems, 2006).

<sup>6</sup> Entry of new generators is generally not the most efficient solution to increase competition. When there is no need for new generation investment, entry, by adding excessive capacity, imposes deadweight losses on the market that can be larger than the gains of increased competition (Green, 1996). Divestiture is in such case the best alternative solution.

also “softer” remedies, such as discouraging incumbents to replace old plants and instead encouraging new entrants to build generation assets, as this is effectively a form of divestiture (no duplication of investment in generation assets).

In addition to such structural remedies, policy makers and regulators have shown interest in behavioral remedies<sup>7</sup> that prevent electricity generators, through the appropriate organization of electricity markets, to be able to use their market power. The wording of EU law suggests that behavioral remedies ought to be the default setting : “Structural remedies should only be imposed either where there is no equally effective behavioural remedy or where any equally effective behavioural remedy would be more burdensome for the undertaking concerned than the structural remedy” (European Commission, 2006a, p.11).

Allaz and Vila (1993) make the theoretical case for the introduction of a forward market as a behavioral remedy that increases competitive pressure.<sup>8</sup> Specifically, analyzing competition in a one-shot game set-up, they show that a forward market lowers the amount of market power producers can exert. The contribution of Allaz and Vila (1993) is important since it has been argued that forward contracts are likely to decrease competition (Lévêque, 2006). Willems et al. (2009), drawing on Allaz and Vila (1993), give the following brief explanation of the effect. In the spot market every producer maximizes his profit given by the profit function  $\pi_i = p[q_i + q_{-i}](q_i - f_i) - c[q_i]$ , where  $q_i$  stands for their own production,  $q_{-i}$  for the production of all other producers, and  $f_i$  for the number of units sold in the forward market. Differentiating this

equation to  $q_i$  and setting it equal to zero yields  $0 = \frac{d\pi_i}{dq_i} \Leftrightarrow -p'[Q](q_i - f_i) = p[Q] - c'[q_i]$ . This

equation can be rewritten<sup>9</sup> as  $\frac{s_i}{E_p^Q} (1 - \frac{f_i}{q_i}) = \frac{p[Q] - c'[q_i]}{p[Q]}$ , where  $s_i$  stands for the market share and

$E_p^Q$  for the price elasticity of demand. We can see from the formula that the markup (the right-hand side of the equation) decreases in  $f_i$ , the number of units sold in the forward market. The more

<sup>7</sup> The European Commission (2006b, p.8) defines a behavioral remedy as “a measure that obliges the concerned undertaking(s) to act in a specific way”.

<sup>8</sup> It has been suggested that a forward market also constitutes a structural remedy. We are agnostic on this definitional issue; after all it is just a label. We note that the EC which defines measures that nudge towards particular actions as behavioral remedies and measures that change the structure of a producer (such as divestiture) as a structural remedy. In general, behavioral remedies are easier to implement than structural remedies.

<sup>9</sup> Multiplying the left side by  $\frac{p[Q] \cdot q_i \cdot Q}{p[Q] \cdot q_i \cdot Q}$  gives  $-p'[Q](q_i - f_i) \cdot \frac{p[Q] \cdot q_i \cdot Q}{p[Q] \cdot q_i \cdot Q} = p[Q] - c'[q_i]$ . Rearranging gives

$$-p'[Q] \frac{Q}{p[Q]} \frac{q_i}{Q} \frac{(q_i - f_i)}{q_i} = \frac{p[Q] - c'[q_i]}{p[Q]}$$

producers sell in the forward market, the closer the outcome in the spot market will be to the Walrasian outcome.

Welfare and consumer surplus thus increase in the number of units sold in the forward market. But do producers have incentives to sell units in the forward market? Allaz and Vila (1993) show that they do. Suppose that only one, privileged, firm could sell in the forward market. In that case this firm has a first mover's advantage. It can, by selling the right number of units in the forward market, reach the Stackelberg equilibrium, which has a higher profit for the privileged firm. Thus, it cannot be a Nash-equilibrium for any firm to sell in the forward market. Consequently, when all firms are entitled to sell in the forward market, they all end up worse off than when none of them had sold. This prisoner's-dilemma type result is standard textbook (e.g., Binmore 2007, chapter 10). Producers earn the highest profit if nobody sells in the forward market, but selling in the forward market is a strictly dominant action for each individual producer. Allaz and Vila (1993) model the competition as a one-shot game. The Nash-equilibrium of a one-shot game and that of a repeated game with predictable ending are theoretically (albeit not necessarily behaviorally) identical. This makes the Allaz and Vila model a fitting theoretical benchmark for an experiment with a fixed, or predictable, number of periods. Of course, in the real world there is no fixed ending of the game (and the Nash-equilibrium of the stage game is one of many equilibria) which might affect the interpretation of the external validity of our results.

In this paper we investigate theoretically and experimentally the effects on competition of introducing forward markets in electricity markets. For relevant parameterizations, we compare the results of introducing a forward market with those of the best alternative remedy: reducing market concentration by divestiture. We do so for competitor numbers that better reflect the market concentration in the old European states than previous literature has done: We also use realistic cost configurations and tease apart competition and asset effect.

We show that, theoretically and behaviorally, the effects of introducing a forward market might be larger than adding one more competitor in markets both with two and three producers. Previously, Brandts, Pezanis-Christou, and Schram (2008) came to the opposite conclusion for the case of three initial competitors.

Whether the theoretical predictions of Allaz and Vila (1993) will materialize in the reality of a dynamic setting such as the EU electricity market has clear policy implications. An affirmative answer would suggest that regulators formulate guidelines for, and promote, the design of effective forward markets.

In the following section we first discuss the experimental design (i.e., the basic parameterizations, treatments, underlying working hypotheses) and experimental procedures as well as related literature. In section 3 we report the results focusing on aggregate quantity, efficiency,



and production efficiency. In section 4 we conclude. The appendices contain robustness tests and instructions.

## 2 Experimental design and procedures

### 2.1 Treatments

We identify the effects of adding one more competitor through divestiture on the one hand and the effects of introducing a forward market on the other, and then compare the effects.

We model the competition of generators in the spot and forward markets using the standard Cournot approach (see for example Borenstein and Bushnell, 1999; LeCoq and Orzen, 2006; Bushnell 2007; Newbery, 2009). The supply-function approach of Klemperer and Meyer (1989) has been argued to be a more accurate approach to model competition in electricity markets. The supply-function approach, however, is more complicated and predicts a wide continuum of equilibria which in turn brings about an equilibrium selection problem (see Devetag and Ortmann, 2007, for a recent review). Wolak and Patrick (2001) provide empirical evidence that dominant generators exert market power by declaring plants to be unavailable, thus shifting the supply curve and suggesting that the Cournot approach is an appropriate modeling choice. In addition, Willems et al. (2009) show that Cournot and supply-function approaches lead to comparable outcomes. In contrast, Green (2004) argues that the Cournot approach does not accurately characterize producer behavior in England and Wales between 1985 and 2000.

Klemperer and Meyer (1989) show that the Cournot equilibrium outcome is the equilibrium with the maximal exertion of market power in the range of supply-function equilibria and hence, arguably, the natural benchmark. Brandt et al. (2008) show that this is also true for configurations with a forward market. The Cournot approach is thus not only relevant and interesting, but can be understood as a necessary first step for additional studies using the supply-function approach.

Table 1 summarizes our treatments and indicates how they compare with earlier studies, namely LeCoq and Orzen (2006) and Brandts et al. (2008), about which more below.

**Table 1: Treatments**

	<b>2 producers</b>	<b>3 producers</b>	<b>4 producers</b>
<b>Without Forward Market</b>	<b>M2<sup>#</sup></b>	<b>M3<sup>*</sup></b>	<b>M4<sup>†</sup></b>
<b>With Forward Market</b>	<b>M2F<sup>#</sup></b>	<b>M3F<sup>*</sup></b>	–
<b>Without Forward Market, zero costs</b>	M2zc <sup>§</sup>	–	–
<b>With Forward Market, zero costs</b>	M2Fzc <sup>§</sup>	–	–

# The treatment is different from the one tested in LeCoq and Orzen (2006) in that our producers face quadratic marginal costs.

† The treatment is different from the one tested in Brandts et al. (2008) in that the market has been created from the market with 3 producers by divestiture, not by entry; producers thus have the same set of assets as those in the market with 3 producers.

§ The treatment is identical to the one tested in LeCoq and Orzen (2006).

\* The treatment is identical to the one tested in Brandts et al. (2008).

A key characteristic is the number of producers in the electricity market. While there is some variance, assuming two producers for markets in the new EU member states<sup>10</sup> and three producers for markets in the old EU member states<sup>11</sup> seems a good approximation.<sup>12</sup>

For the NMS-12 we thus compare outcomes in markets with two producers and without a forward market (M2) with outcomes in such markets with a forward market (M2F). We also compare the difference in outcomes with the difference in outcomes of markets with two (M2) and three producers (M3), when for the latter we add one more producer by means of divestiture. In other words, we compare the differences of  $M2F - M2$  and  $M3 - M2$ . The markets M2zc and M2Fzc are treatments to allow comparison of our results with the experimental results of LeCoq and Orzen (2006).

For the EU-15 we compare outcomes in markets with three producers and without a forward market (M3) with outcomes in such markets with a forward market (M3F). We also compare the difference in outcomes with the difference in outcomes of markets with three (M3) and four producers (M4), when for the latter we add one more producer by means of divestiture. In other words, we compare the differences of  $M3F - M3$  and  $M4 - M3$ .

## 2.2 Earlier experiments

LeCoq and Orzen (2006) conducted experiments in markets with two producers with and without a forward market and compared the outcomes with those in a market with four producers (with and without a forward market); importantly, their producers faced zero production costs. In line with earlier experiments, such as Huck et al. (2004), LeCoq and Orzen (2006) found that producers competed less (more) than predicted with two (four) producers. A forward market had a positive effect, but weaker than expected. Adding two more producers increased output significantly more than introducing a forward market.

LeCoq and Orzen (2006) consider the effects of a forward market in a market with two (and four) producers. While speaking possibly to the reality of electricity markets in the NMS-12 countries, the number of relevant competitors tends to be three for EU-15 countries. Moreover, the

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<sup>10</sup> The new EU member states are the states that acceded to the EU in or after 2004. With the exception of Cyprus and Malta they are all post-communist countries: Bulgaria (BG), Cyprus (CY), the Czech Republic (CZ), Estonia (EST), Hungary (H), Lithuania (LT), Latvia (LV), Malta (M), Poland (PL), Romania (RO), Slovakia (SK), and Slovenia (SLO).

<sup>11</sup> The old EU member states are the states that joined the EU before 2004. These are: Austria (A), Belgium (B), England (UK), Germany (D), Denmark (DK), Spain (E), France (F), Finland (FIN), Greece (GR), Italy (I), Ireland (IRL), Luxembourg (L), the Netherlands (NL), Portugal (P), Sweden (S).

<sup>12</sup> The average Hirsch-Herfindahl Index (HHI) for the old (West European) EU members in 2006 was equal to 3786, which is close to the case where three symmetrical firms compete ( $HHI=3333$ ). The new (Central and East European) EU members had in 2006 a HHI equal to 5558, which is closer to the case where two symmetrical firms compete ( $HHI=5000$ ) (Van Koten and Ortmann, 2008).

assumption that producers have zero marginal costs is unrealistic for all scenarios. In our experiment, producers therefore face, more realistically (e.g., Newbery, 2002) and in line with Brandts et al. (2008), quadratic marginal costs.

Brandts et al. (2008) conducted experiments in markets with three producers with and without a forward market and compared the outcomes with those in a market with four producers (without a forward market). Producers had quadratic marginal costs. Brandts et al. (2008) find that a forward market significantly increases the quantity supplied, but that entry of a new generator increases the quantity supplied significantly more than the addition of a forward market.

Brandts et al. (2008) confound two effects in their study: the competition effect<sup>13</sup> and the asset effect. The competition effect is brought about by an additional market participant; this makes the market more competitive and results in lower prices and a larger total number of units supplied. The asset effect is brought about by the additional production assets that are built and paid by a new entrant. Because Brandts et al. (2008) consider the entrance of a new generator, their treatment combines the competition and the asset effect: entrance increases competition, but also the aggregate size of production assets in the market, which reduces the aggregate cost and thus gives an extra incentive to increase production. Thus, assuming efficient production, any given level of aggregate production (the production of all producers together) is produced more cheaply in the market with four producers than in the market with three producers. We conjecture that the asset effect confound led to an overestimation of the effects of adding one more competitor in their study. Moreover, the welfare effects Brandts et al. (2008) reports are not conclusive, as they do not incorporate the very high costs of the increase in the asset base (the cost of building extra production plants).<sup>14</sup>

We therefore focus on the effect of divestiture as a benchmark for the effects of introducing a forward market, thus eliminating the asset effect confound and insulating the competition effect. To allow for comparisons, we drew (to the extent possible) on Brandts et al. (2008) and on LeCoq and Orzen (2006) to parameterize our experiment.<sup>15</sup>

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<sup>13</sup> Brandts et al. (2008) call this effect the “number effect”.

<sup>14</sup> Building electricity generation is very costly, and when the problem is a lack of competition but not a shortage of electricity production capacity, entrance leads to a wasteful duplication of assets (Green, 1996).

<sup>15</sup> Another paper addressing forward markets is Ferreira, Kujal and Rassenti (2009). Ferreira et al. (2009) find that forward markets have a positive effect on the aggregate production for inexperienced subjects, but no - or a negative - effect with experienced subjects. Producers in their experiment face linear rather than (as in our set-up) quadratic marginal costs. Also, subjects had a mere 30 seconds to make a decision and no extra time to review the results, whereas the subjects in our study had 60 seconds to make a decision and another 60 seconds to review the result of that decision. Moreover, in the experiment of Ferreira et al. (2009) the treatments with experienced subjects were different from the treatments with inexperienced subjects in important details. For example, the experienced subjects did not have production costs and the demand function they faced was different. In addition, due to random matching, the experiment draws on very few independent data points (1 or 2).

We found the result that experience decreases the competitive effect of a forward market nonetheless interesting and, as a robustness test, ran additional sessions with experienced subjects. Our subjects were experienced in the sense that they had participated in the experiment earlier. Our experienced subjects were assigned to exactly the same

### 2.3 Demand and supply

As in Brandts et al. (2008), the demand schedule is  $p(Q) = \text{Max}(0, 2000 - 27Q)$ ,  $Q \geq 0$ . Also as in Brandts et al. (2008), we chose to program the demand side rather than have it enacted by experimental participants. This might reduce demand uncertainty which in turn is likely to influence (the speed of) convergence in our market. We believe that this choice does not interact with treatment effects in a significant manner.

For some treatments we model generators as having quadratic marginal costs. Marginal costs of producing electricity usually have a hockey-stick shape, i.e., they are flat with a sharp increase when capacity constraints become binding (Newbery, 2002). We consider marginal quadratic costs to be a reasonable approximation to the real cost curves of electricity generators.

To be able to compare our results with those of Brandts et al. (2008), we also use the same specification of the costs for markets with three producers, abbreviated by M3 for the market without forward market and by M3F for the market with forward market. Brandts et al. (2008) set the marginal cost of producing the  $i^{\text{th}}$  unit for a producer equal to  $mc_3(q) = 2x^2$ , cumulative costs

can thus be calculated as  $c_3(q) = \sum_{x=1}^q 2x^2 = \frac{2}{3}x^3 + x^2 + \frac{1}{3}x$ .

The market with four producers, M4, is created from the market with three producers, M3, by divestiture; each of the three producers divests  $\frac{1}{4}^{\text{th}}$  of their assets, and these three sets of assets are used to create a fourth, identical producer. The markets with two producers, M2 and M2F, are created from the market with three producers, M3, by reversing the divestiture process (merger): one of the producers is split in half and its assets are merged to the two remaining producers to create two larger, identical, producers. With the cost function of a producer in M3 given, the cost

functions of producers in M2 and M4 can be calculated:  $c_2[\mathbf{y}] = \frac{8\mathbf{y}^3}{27} + \frac{2\mathbf{y}^2}{3} + \frac{\mathbf{y}}{3}$ , and

$$c_4[\mathbf{y}] = \frac{32}{27}\mathbf{y}^3 + \frac{4}{3}\mathbf{y}^2 + \frac{\mathbf{y}}{3}.^{16}$$

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treatment as the one in which they had participated earlier (with the exception for experienced subjects in treatment M2, where some subjects had participated earlier in M3 or M4). Overall we do not find that experienced subjects have a lower aggregate production than inexperienced subjects. On the contrary, we find indications that experienced subjects supply a slightly higher production, which is in line with the experimental literature on the effect of experience on public good provision (Ledyard, 1995). We report the detailed results as a robustness test in the Appendix.

<sup>16</sup> If all the assets in M3 would be merged into one single firm (M1), then this single firm would minimize its costs by dividing production equally over the three plants. The total costs of the firm in M1 would thus be  $c_1[\mathbf{y}] = 3 \cdot c_3[\frac{\mathbf{y}}{3}]$ . Likewise, if we started with M2 and merged the assets of the two firms into one single firm, then the cost function of the single firm in M1 would be given by  $c_1[\mathbf{y}] = 2 \cdot c_2[\frac{\mathbf{y}}{2}]$ . Now we can derive the cost function for a firm in M2: from  $2 \cdot c_2[\frac{\mathbf{y}}{2}] = c_1[\mathbf{y}] = 3 \cdot c_3[\frac{\mathbf{y}}{3}]$  follows that  $c_2[\mathbf{y}] = \frac{3}{2} \cdot c_3[\frac{2}{3} \cdot \mathbf{y}] = \frac{8}{27}\mathbf{y}^3 + \frac{2}{3}\mathbf{y}^2 + \frac{1}{3}\mathbf{y}$ . In the same manner we can derive

The electricity generation asset base is the same for all three markets (M2, M3, and M4). Therefore, when generators make identical choices and the aggregate production is equal over different markets, the aggregate costs must also be equal.

Table 2 summarizes the production costs for each generator in the market with two (M2), three (M3) and four (M4) generators, and identifies aggregate production in one market being equal to that in another market by bold numbers and shaded cells. For example, the aggregate production in M2 (M4) is equal to that in M3 when the total number of units can be divided both by two (four) and three.

To not unduly add to our subjects' cognitive load, we presented costs that were rounded according to the following rounding rules:

- All numbers smaller than 100 were rounded to the nearest integer number.
- When a number was larger than 100, it was rounded to the nearest 5-fold.
- When a number was larger than 1000, it was rounded to the nearest 10-fold.
- When a number was larger than 10000, it was rounded to the nearest 50-fold.

As a result of these rounding rules, some of the aggregate total costs in Table 2 are different. The discrepancy is small, however; on average the absolute discrepancies is 0.12%. For the "rounded numbers" version of Table 2, see Table A1 in the Appendix.

The numbers we obtained after this rounding procedure were also the numbers we use to calculate the theoretical predictions.<sup>17</sup>

**Table 2:<sup>18</sup> Overview of aggregate cost of producing**

Market with two producers (after merger)			Market with three producers (original market)					Market with four producers (after divestment)						
Each Producer		Aggregate		Each Producer			Aggregate		Each Producer			Aggregate		
Units produced by each producer	Marginal Costs	Total Costs	Total Production	Total Costs	Units produced by each producer	Marginal Costs	Total Costs	Total Production	Total Costs	Units produced by each producer	Marginal Costs	Total Costs	Total Production	Total Costs
N	MC	TC	2*N	2*TC	N	MC	TC	3*N	3*TC	N	MC	TC	4*N	4*TC
0	0	0	<b>0</b>	<b>0</b>	0	0	0	<b>0</b>	<b>0</b>	0	0	0	<b>0</b>	<b>0</b>
1	1	1	2	2	1	2	2	3	6					
2	5	6	<b>4</b>	<b>11</b>						1	3	3	<b>4</b>	<b>11</b>
3	9	15	<b>6</b>	<b>30</b>	2	8	10	<b>6</b>	<b>30</b>					
4	16	31	<b>8</b>	<b>62</b>						2	12	15	<b>8</b>	<b>62</b>
5	24	55	10	111	3	18	28	9	84					

the cost function for a firm in M4:  $c_4[y] = \frac{3}{4} \cdot c_3[\frac{4}{3} \cdot y] = \frac{32}{27} y^3 + \frac{4}{3} y^2 + \frac{1}{3} y$ . Notice that for marginal costs holds the equality:  $c_2'[\frac{3}{2} y] = c_3'[y] = c_4'[\frac{3}{4} y]$ . Conforming to intuition, the marginal cost of a firm in M3 thus increases faster (slower) than in M2 (M4).

<sup>17</sup> Using the exact numbers gives virtually identical theoretical predictions.

<sup>18</sup> Numbers have been rounded to the nearest whole number.

6	35	90	12	180	4	32	60	12	180	3	30	45	12	180
7	47	137	14	273	5	50	110	15	330					
8	60	197	16	394						4	54	99	16	394
9	76	273	18	546	6	72	182	18	546					
10	93	366	20	733						5	84	183	20	733
11	113	479	22	957	7	98	280	21	840					
12	133	612	24	1224	8	128	408	24	1224	6	123	306	24	1224
13	156	768	26	1536	9	162	570	27	1710					
14	180	948	28	1897						7	168	474	28	1897
15	207	1155	30	2310	10	200	770	30	2310					
16	235	1390	32	2779						8	221	695	32	2779
17	264	1654	34	3308	11	242	1012	33	3036					
18	296	1950	36	3900	12	288	1300	36	3900	9	280	975	36	3900
19	329	2279	38	4559	13	338	1638	39	4914					
20	365	2644	40	5287						10	347	1322	40	5287
21	401	3045	42	6090	14	392	2030	42	6090					
22	440	3485	44	6970						11	420	1742	44	6970
23	480	3965	46	7931	15	450	2480	45	7440					
24	523	4488	48	8976	16	512	2992	48	8976	12	502	2244	48	8976
25	567	5055	50	10109	17	578	3570	51	10710					
26	612	5667	52	11334						13	590	2834	52	11334
27	660	6327	54	12654	18	648	4218	54	12654					
28	709	7036	56	14073						14	684	3518	56	14073
29	761	7797	58	15593	19	722	4940	57	14820					
30	813	8610	60	17220	20	800	5740	60	17220	15	787	4305	60	17220
31	868	9478	62	18956	21	882	6622	63	19866					
32	924	10402	64	20805						16	896	5201	64	20805
33	983	11385	66	22770	22	968	7590	66	22770					
34	1043	12428	68	24855						17	1013	6214	68	24855
35	1104	13532	70	27064	23	1058	8648	69	25944					
36	1168	14700	72	29400	24	1152	9800	72	29400	18	1136	7350	72	29400
37	1233	15933	74	31867	25	1250	11050	75	33150					
38	1301	17234	76	34467						19	1267	8617	76	34467
39	1369	18603	78	37206	26	1352	12402	78	37206					
40	1440	20043	80	40086						20	1405	10022	80	40086
41	1512	21555	82	43111	27	1458	13860	81	41580					
42	1587	23142	84	46284	28	1568	15428	84	46284	21	1549	11571	84	46284
43	1663	24805	86	49609	29	1682	17110	87	51330					
44	1740	26545	88	53090						22	1702	13273	88	53090
45	1820	28365	90	56730	30	1800	18910	90	56730					
46	1901	30266	92	60533						23	1860	15133	92	60532
47	1985	32251	94	64501	31	1922	20832	93	62496					
48	2069	34320	96	68640	32	2048	22880	96	68640	24	2027	17160	96	68640

## 2.4 Theoretical Predictions and Hypotheses

With demand given and the cost function defined, the profit function is given by

$\pi_{i,MS} = p[q_i + q_{-i}](q_i - f_i) - c_{MS}[q_i]$  for each of the market sizes  $MS \in (2,3,4)$ , where the cost

functions are defined as above by  $c_2[y] = \frac{8y^3}{27} + \frac{2y^2}{3} + \frac{y}{3}$ ,  $c_3(q) = \sum_{x=1}^q 2x^2 = \frac{2}{3}x^3 + x^2 + \frac{1}{3}x$ , and

$c_4[\mathbf{y}] = \frac{32}{27}y^3 + \frac{4}{3}y^2 + \frac{y}{3}$ . We can now determine the Nash-equilibria for each of the treatments.

Table 3 shows the theoretical predictions for our treatments M2, M2F, M3, M3F, and M4.<sup>19</sup> The prefix NE stands for Nash-equilibrium, Walras for the efficient solution (the outcome that maximizes the total surplus),<sup>20</sup> and JPM for Joint Profit Maximization (the monopoly solution).<sup>21</sup>

**Table 3: Theoretical predictions for electricity markets**

	NE M2	NE M2F		NE M3	NE M3F	NE M4	Walras (n=2)	Walras (n=3)	Walras (n=4)	JPM (n=2)	JPM (n=3)	JPM (n=4)
$q_{ii}^f$	–	2	11	–	5	–	–	–	–	–	–	–
$q_{ii}$	20	20	22	14/15 <sup>22</sup>	15	11	25/26 <sup>23</sup>	17	13	16	11	8
$q_t$	40	40	44	43	45	44	51	51	52	32	33	32
$p_t$	920	920	812	839	785	812	623	623	596	1136	1109	1136
<i>Prod. S.</i>	31520	31520	28768	29537	27885	28768	21053	21063	19672	33572	33567	33572
<i>Cons. S.</i>	21060	21060	25542	24381	26730	25542	34425	34425	35802	13392	14256	13392
<i>Total S.</i>	52580	52580	54310	53918	54615	54310	55478	55488	55474	46964	47823	46964
<i>Eff. (%)</i>	94.8	94.8	97.9	97.2	98.4	97.9	100	100	100	84.7	86.2	84.7

<sup>19</sup> The Nash equilibria have been numerically determined with Mathematica programs. The set of programs can be downloaded as a RAR file named “Nash-Equilibria with Forward Markets.RAR”, at <https://sites.google.com/site/slvstrnl/ElectricityMarketsExperiment>. The predictions are based on the cost functions with numbers rounded according to the rounding procedure described above. Predictions based on the continuous cost functions are, except for the M2F condition, mostly identical: the chosen quantities are identical, and the difference in total surplus is lower than 0.02%. In the M2F condition the chosen quantities in the low Nash-equilibrium are lower when using the continuous functions – 40 instead of 42. As a result the difference in total surplus is relatively high: 1.8%.

<sup>20</sup> We define efficiency, following Brandts et al. (2008), as the joint consumer and producer surplus realized in the experiment divided by the maximum joint consumer and producer surplus (the Walrasian level of joint surplus).

<sup>21</sup> The markets JPM (n=3), JPM (n=4), NE C3.0, NE C3.2, Walras (n=3), Walras (n=4) and NE C4.0 in this experiment are identical to those in Brandts et al. (2008), and our predictions are almost identical to the ones reported in their paper. Key differences are: Using the functions without a rounding procedure, we find that for the Nash-equilibrium with three producers (M3) the price is equal to 839 rather than 866, as reported in Brandts et al. (2008). We find that for the Nash-equilibrium with four generators (M4), the price is equal to 677 rather than 704. Also, the producer surplus of M4 is equal to 27635 rather than 27638. For the welfare maximizing outcome with four generators, Walras (n=4), we find that all three generators produce 14 units and one of them 15 units, instead of all of the generators producing 14 units. Total welfare is therefore 60799 and not 60788. For the monopoly case with four generators, JPM (n=4), two generators produce 9 units and two 8 units, instead of all of them 8 units. As a result the producer surplus is higher, 34832 instead of 34728, the consumer surplus is lower, 15147 instead of 17010, and efficiency is lower, 82.2% instead of 85.1%.

For the Nash-equilibrium with three producers and a forward market (M3F), we find a unique symmetrical Nash-equilibrium in pure strategies where each producer sells 5 units in the forward market, and 10 additional units in the spot market. This is different from Brandts et al. (2008), who for the treatment with the forward market (M3F) consider partially mixed strategies (for the choice of additional units) and find an equilibrium where each producer sells 6 units in the forward market, and an additional 9 with probability .944 and 10 with probability 0.056. As we find a unique symmetric Nash-equilibrium in pure strategies, we do not follow Brandts et al. (2008) in broadening the equilibrium concept for one treatment case (no mixed strategies are considered for the other treatments). In any case, the total (expected) production by all three producers we find and the one reported by Brandts et al. (2008) are the same – 45 units.

<sup>22</sup> One generator produces 15 units, the other two 14 units.

<sup>23</sup> One generator produces 26 units, the other two 25 units.

The theoretical predictions give us, for the particular parameterizations chosen, an indication of the effect on aggregate production and efficiency of introducing a forward market or adding one more competitor. For markets with three producers, both introducing a forward market and adding one more competitor increases aggregate production, but introducing a forward market increases aggregate production more. For markets with two producers, adding one more competitor increases aggregate production. Introducing a forward market increases aggregate production only if the higher Nash-equilibrium is realized. In fact, aggregate production in that case is increased more than in the case of one more competitor. Using  $q(x)$  to denote aggregate production in market structure  $x$ <sup>24</sup>, we thus conjecture that the remedies can be ranked as follows:  $q(\text{M3F}) > q(\text{M4}) > q(\text{M3})$ . Likewise, both remedies also increase efficiency, but introducing a forward market again is predicted to increase efficiency the most. Using  $\Omega(x)$  to denote efficiency in market structure  $x$ , we thus conjecture that the remedies can be ranked as follows:  $\Omega(\text{M3F}) > \Omega(\text{M4}) > \Omega(\text{M3})$ .

For markets with two producers, both introducing a forward market and adding one more competitor increases aggregate production, but the existence of two welfare-rankable Nash-equilibria makes it impossible to rank the remedies. We conjecture that the remedies can be ranked as follows:  $q(\text{M2F}) > q(\text{M2})$ ,  $q(\text{M3}) > q(\text{M2})$ , and  $q(\text{M2F}) = q(\text{M3})$ . Moreover, the theoretical results, assuming the lower-ranked Nash-equilibrium will occur at least with some probability,<sup>25</sup> suggest that the effect of introducing a forward market is not as large as adding two more competitors; we thus conjecture  $q(\text{M4}) > q(\text{M2F})$ . Both remedies also increase efficiency but again they cannot be ranked. We conjecture that:  $\Omega(\text{M2F}) > \Omega(\text{M2})$ ,  $\Omega(\text{M3}) > \Omega(\text{M2})$ ,  $\Omega(\text{M2F}) = \Omega(\text{M3})$ , and  $\Omega(4) > \Omega(\text{M2F})$ .

**Table 4: Hypotheses**

<b>Hq.1 (Quantity)</b>	<b>HΩ.1 (Efficiency)</b>	<b>HΦ.1 (Production Efficiency)</b>
- $q(\text{M3F}) > q(\text{M4}) > q(\text{M3})$	- $\Omega(\text{M3F}) > \Omega(\text{M4}) > \Omega(\text{M3})$	- $\Phi(\text{M3F}) = \Phi(\text{M3})$
		- $\Phi(\text{M4}) < \Phi(\text{M3})$
<b>Hq.2 (Quantity)</b>	<b>HΩ.2 (Efficiency)</b>	<b>HΦ.2 (Production Efficiency)</b>
- $q(\text{M2F}) > q(\text{M2})$	- $\Omega(\text{M2F}) > \Omega(\text{M2})$	- $\Phi(\text{M2F}) = \Phi(\text{M2})$
- $q(\text{M3}) > q(\text{M2})$	- $\Omega(\text{M3}) > \Omega(\text{M2})$	- $\Phi(\text{M3}) < \Phi(\text{M2})$
- $q(\text{M2F}) = q(\text{M3})$	- $\Omega(\text{M2F}) = \Omega(\text{M3})$	
<b>Hq.3 (Quantity)</b>	<b>HΩ.3 (Efficiency)</b>	
- $q(\text{M4}) > q(\text{M2F})$	- $\Omega(\text{M4}) > \Omega(\text{M2F})$	

<sup>24</sup> To facilitate comparisons with related literature, we use the same notation as Brandts et al. (2008). Also parts of our presentation have been inspired by Brandts et al. (2008).

<sup>25</sup> M2F has two Nash equilibria outcome for the aggregate production: a low one of 40 and a high one of 44. The high Nash equilibrium in M2F is equal to the Nash equilibrium in M4. The aggregate production in M2F can thus be expected to be lower than that in M4, as long as the lower Nash equilibrium of 40 occurs with at least some probability.



We also test for effects on production efficiency. We define production efficiency, following Brandts et al. (2008), as actual producer surplus divided by producer surplus had production taken place in the most efficient manner. As marginal costs are quadratic, production is fully efficient only if the aggregate production is evenly distributed over the producers. Like Brandts et al. (2008) we assume that more producers in a market should make it more difficult to achieve an even distribution, but that introducing a forward market should not have an effect. We thus conjecture  $\Phi(M4) < \Phi(M3) < \Phi(M2)$ ,  $\Phi(M3F) = \Phi(M3)$ , and  $\Phi(M2F) = \Phi(M2)$ . Table 4 summarizes our hypotheses.

## 2.5 Experimental procedures

The experimental sessions were conducted in October 2009, December 2009, and April 2010 at CERGE-EI in Prague.<sup>26</sup> Our subjects were students at Charles University or at the University of Economics, both located in Prague. A total of 198 students participated. The sessions with a forward market lasted about 2 hours, the sessions without forward market lasted about 90 minutes. The subjects earned on average 382 Czech Koruna per hour including a show-up fee of 100 Korunas, meaning that subjects earned on average 640 Korunas.<sup>27</sup> The minimum earning was 330 and the maximum earning was 1080 Korunas, indicating that our experiment was well-incentivized on the margin. At the beginning of each session, the English instructions were read to subjects by the experimenter (Van Koten).

The market simulation was programmed in Z-Tree (Fischbacher, 2007). The demand schedule was pre-programmed. Participants took on the roles of producers and sellers. They were not shown the demand schedule but were given on screen, and as a printout, a payoff table.

In the treatments with a forward market every round has two periods: the first period for the forward market and the second period for the spot market. In the first period, producers decide how many units to produce and to sell in the forward market. Producers sell units to traders. The units that producers sell are promises to produce and to deliver units to traders in the second period (in the spot market). The units that are sold in the forward market are thus produced later, in the second period. To help producers see the effect of their actions on their profits, we communicated to them the cost of their selling decision in the forward market and the resulting profit.

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<sup>26</sup> We obtained in October 2009 four data points for each treatment, in December 2009 four data points for M2zc, M2Fzc, M2F, M2, M3, and three data points for treatments M3F and M4, and in April 2010 three data points for M2zc, M2Fzc, M2F, M2, M3, and four data points for treatments M3F and M4. The original plan was to obtain four data points for all treatments also in December 2009. Unusual numbers of no-shows for treatments M3F and M4 derailed that plan. Several pilot sessions were run during the summer of 2009. None of the subjects in the pilot (mostly CERGE-EI students) participated in the regular sessions.

<sup>27</sup> This equaled about €26 (and about €36 at official purchasing power parity, and even more on student-specific purchasing power parity).

In the forward market two pre-programmed traders compete in prices for the total number of units that are offered.<sup>28</sup> We do not explicitly mention the existence of traders to the subjects. Because traders act rationally, their actions define a demand schedule, and we present this schedule to our producers.<sup>29</sup> The trader that offers the highest price per unit wins all units. When they offer the same price – which they do in equilibrium – a winner is drawn at random. As the pre-programmed traders are rational and compete in prices, they can predict the Nash-equilibrium spot price and offer this price for the units offered in the forward market. The pre-programmed traders do not observe the number of units offered by each producer, only the total number of units. They assume that the producer offers an equal number of units in the forward market.<sup>30</sup> Using this assumption, the traders are programmed to predict, conditional on the total number of units offered in the forward market,  $q^f$ , the Nash-equilibrium total production in the spot market:  $q^{NE}(q^f)$ .<sup>31</sup> By substituting the predicted total production in the spot market in the demand schedule, the traders predict the Nash-equilibrium price in the spot market:  $p(q^{NE}(q^f))$ . As traders offer the Nash-equilibrium price for all units,  $p(q^{NE}(q^f))$  defines the demand schedule in the forward market. This forward market demand schedule is presented to producers in the first period of each round, so they can use this information when deciding how many units to offer for the forward market. At the end of the period, all producers are paid the number of units they produced in the forward market times the price per unit minus the production cost. Appendix A3 shows, conditional on the total production in the forward market (stage A), the predicted aggregate production and price in the spot market.

In the second period of each round, producers decide how many units to produce and sell in the spot market. The pre-programmed traders sell all the units they bought. The price per unit is determined by substituting the number of units sold by all producers in the forward and spot market together for  $Q$  in the demand schedule  $p(Q) = \text{Max}(0, 2000 - 27Q)$ . All producers are paid the number of units they produced in the spot market times the price per unit minus the production cost.

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<sup>28</sup> In our experiment traders are not represented by experimental subjects, but they are, as in LeCoq and Orzen (2006), pre-programmed. The manner in which traders are represented in the experiment should not significantly affect outcomes, as traders are middlemen (between producers in the forward market and end demand in the spot market). Earlier experimental evidence indicates that the presence of strategically-acting middlemen generally does not alter allocations and that the profit of middlemen converges to zero (Plott and Uhl, 1981). Brandts et al. (2008) use experimental subjects as traders and find indeed that a trader earns only a small portion (about 8%) of the profits that a producer earns.

<sup>29</sup> The full consolidated instructions can be downloaded at <https://sites.google.com/site/slvstrnl/ElectricityMarketsExperiment>.

<sup>30</sup> When the total of units sold forward is not divisible by the number of producers in the market, the trader assumes that the numbers of units sold forward by each producer are as close as possible. For example, when the total number of units sold forward is 14 in M3, the trader assumes that two producers sold 5 units and one producer 4 units. Violation of this assumption affects the prediction only minimally.

<sup>31</sup> This procedure is virtually identical to the one used in LeCoq and Orzen (2006).

### 3. Results

We have 11 statistically independent data points for all treatments (each data point below we call “a group” consisting of the aggregate of sellers in a particular treatment); since each participant took part in one experimental session, data points are also statistically independent across treatments. None of the participants went bankrupt. Each treatment consisted of 24 rounds. For our statistical tests, we use only the last 12 rounds of the data, as the experiment is complicated and we know – for example, from relatively easy auction experiments – that subjects need several rounds of trading to become familiar with the laboratory environment before they react to the embedded incentives (Hertwig and Ortmann, 2001). Following LeCoq and Orzen (2006), we test for disparity with the Nash-equilibrium predictions using two-sided Wilcoxon one-sample signed-rank tests (two-sided signed-rank tests), unless indicated otherwise. For comparison between the averages of the treatment in our experiment, we use, following Brandts et al. (2008), F-tests based on an OLS regression of the dependent variable on the 5 treatment dummies, M2, M2F, M3, M3F, and M4, without a constant (F-tests). The error terms are adjusted for clustered data by using the robust Huber/White/Sandwich estimator (Froot, 1989). To compare three ordered inequalities, we also run, following Brandts et al. (2008), a Jonckheere test, which makes no distributional assumptions. In addition, we ran robustness tests using, as did LeCoq and Orzen (2006), Wilcoxon rank-sum tests (rank-sum tests). These tests confirmed most of the results presented here. The results of these tests may be found in Appendix A2.

#### 3.1. Aggregate Quantity

Figure 1 shows the evolution of total (aggregate) quantities sold per period, averaged over treatment groups. Treatments with two traders are represented by circles, with three traders by triangles, and with four traders by squares. The treatments without forward markets are represented by open circles, triangles or squares, the treatments with forward markets by filled circles or triangles.

The volume in all treatments starts out rather low<sup>32</sup> but trade volume moves quickly into the direction of the Nash-equilibrium. Between rounds 8 and 12 behavior has stabilized.

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<sup>32</sup> It is likely that these trajectories are anchored by the examples in the instructions; in the examples we used low numbers to facilitate understanding of the basic relationships.

**Figure 1 : Aggregate production**

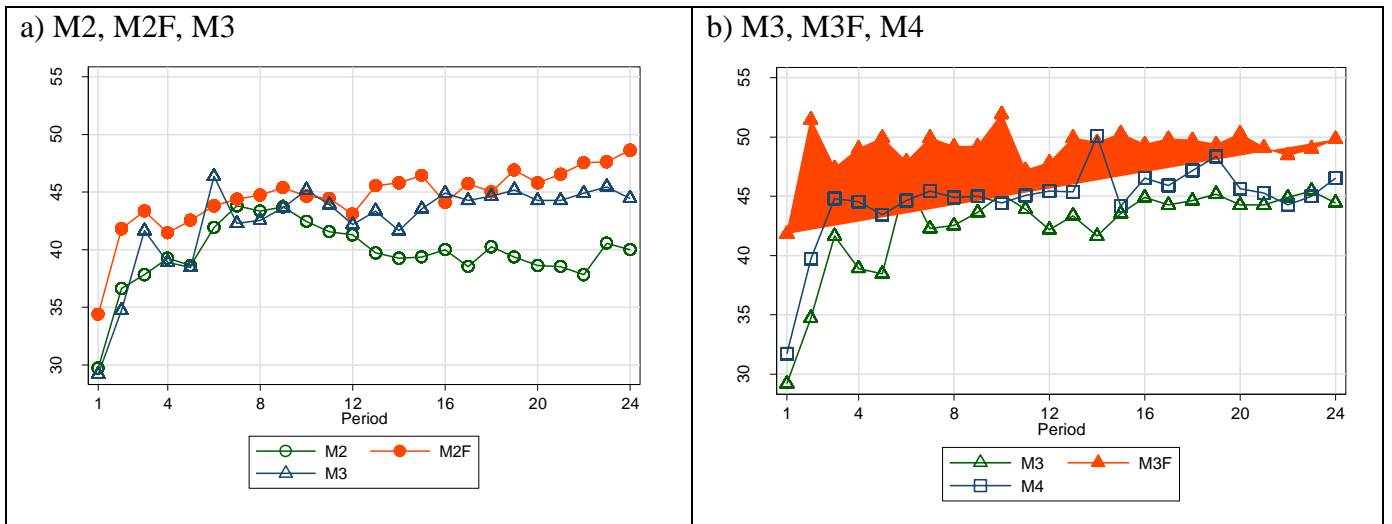


Table 5 shows the overall average aggregate production per treatment group, with the standard error in parenthesis.<sup>33</sup> The row below gives the size of the observed aggregated quantity relative to the Nash-equilibrium prediction in percentages.

**Table 5: Production averages in the last 12 rounds**

	M2	M2F	M3	M3F	M4
Average production	39.3 (1.52)	46.3 (2.06)	44.2 (1.22)	49.6 (0.61)	46.2 (0.98)
% of NE prediction	98.7%	116 % / 105% <sup>34</sup>	102.9%	110.1%	105.0%
Number of observations	N=11	N=11	N=11	N=11	N=11
% of NE prediction - earlier studies <sup>35</sup>	93,2%, <i>LeCoq and Orzen (2006)</i> 92,7%, <i>Huck et al. (2004)</i>	93,8%, <i>LeCoq and Orzen (2006)</i>	102.7%, <i>Huck et al. (2004)</i> 98.9%, <i>Brandts et al. (2008)</i>	103.6%, <i>Brandts et al. (2008)</i>	113.7%, <i>LeCoq and Orzen (2006)</i> 102.8%, <i>Brandts et al. (2008)</i> 102.9%, <i>Huck et al. (2004)</i>

Note that in the M2 and M2F conditions the standard error is relatively high. Of the treatments without forward markets, M2 and M3 are not significantly different from the Nash-equilibrium predictions (two-sided signed-rank test, both p-values > 0.32), while M4 is significantly larger (p-value=0.068). Of the treatments with a forward market, the production in M3F is significantly

<sup>33</sup> The standard error is computed based on the values of the averages for each group over the last 12 rounds.

<sup>34</sup> The first number gives the percentage of efficiency relative to the low production Nash-equilibrium, the second number relative to the high production Nash-equilibrium.

<sup>35</sup> The averages by Huck et al. (2004) reported here are based on their meta-analysis of 19 experiments with Cournot competition. A Wilcoxon signed-rank test indicates that our results are not significantly different from their results (p-values for M2, M3 and M4 are 0.155, 0.657 and 0.534 respectively). Compared with Brandts et al. (2008), the production is significantly higher in condition M3F (p < 0.006) and not significantly different in the conditions M3 (p-value=0.213) and M4 (p-value=0.534). Compared with LeCoq and Orzen (2006), production is significantly higher in conditions M2F (p-value=0.010 for the low and p-value=0.033 for the high Nash-equilibrium) and M4 (p-value=0.010) and not significantly different in condition M2 (p-value= 0.182). For comparison, we also ran treatments with zero production costs, M2zc and M2Fzc. In these treatments the average production is 83% of the Nash-equilibrium prediction, which is significantly lower than what LeCoq and Orzen (2006) found (both p-values < 0.041). The results of these tests may be found in Appendix A2.

higher than the Nash-equilibrium (p-values = 0.004) and production in M2F is significantly higher than the low Nash-equilibrium (p-value=0.021), but is not significantly different from the high Nash-equilibrium (p-value=0.248).

Without a forward market, when the number of competitors is equal to two (three or four), production tends to be smaller (larger) than the Nash-equilibrium, which is in line with earlier findings (LeCoq and Orzen, 2006; Huck, Normann, and Oechssler, 2004). We see no evidence for long-lasting collusion; indeed the data suggest the opposite. A regression of aggregate production on the period of the experiment shows a significant upwards slope, suggesting that over time, as subjects become more experienced with the task, they become less likely to collude.

**Table 6: Effects of one more competitor and forward market on quantities, Hq.1, Hq.2, and Hq.3**

	OLS regression, with correction for clustering on group level, followed by an one-sided F-test on equality of the coefficients			Jonckheere test
Hq.1 - Markets with 3 producers	$q(M3F) > q(M3)^{***}$ (p<0.001)	$q(M4) > q(M3)$ (p=0.105)	$q(M3F) > q(M4)^{***}$ (p=0.002)	$q(M3F) \geq q(M4) \geq q(M3)$ , with at least one of the inequalities being strict p-value = 0.0000
	N= 792	N= 924	N= 924	N= 1320
Hq.2 - Markets with 2 producers	$q(M2F) > q(M2)^{***}$ (p=0.003)	$q(M3) > q(M2)^{**}$ (p= 0.006)	$q(M2F) = q(M3)$ (p=0.374)	$q(M2F) \geq q(M3) \geq q(M2)$ , with at least one of the inequalities being strict*** p-value = 0.0000.
Number of observations	N= 528	N= 660	N= 660	N= 924
Hq.3	$q(M4) > q(M2F)$ (p=0.521)			
	N= 792			

Table 6 presents the test for our hypothesis using F-tests based on an OLS regression and Jonckheere tests.<sup>36</sup>

<sup>36</sup> As a robustness test we also compared the averages for the groups using a two-sample Wilcoxon rank-sum (Mann-Whitney) test. The hypotheses accepted (rejected) are the same, except for Hypothesis 2.b (which becomes insignificant) and Hypothesis 3.c (which becomes significant). See the Appendix for a detailed analysis.

**Results testing Hypothesis q.1: In markets with 3 competitors, introducing a forward market increases production, and the effect is stronger than adding one more competitor,  $q(M3F) > q(M3)$ , and  $q(M3F) > q(M4)$ .**

We find partial support for Hypothesis q.1:

- $q(M3F) \leq q(M3)$  is REJECTED in favor of  $q(M3F) > q(M3)$ , p-value<0.001.
- $q(M4) \leq q(M3)$  is NOT rejected in favor of  $q(M4) > q(M3)$  , p-value=0.105.
- $q(M3F) \leq q(M4)$  is REJECTED in favor of  $q(M3F) > q(M4)$  , p-value=0.002 .
- $q(M3F) = q(M4) = q(M3)$  is REJECTED in favor of  $q(M3F) \geq q(M4) \geq q(M3)$ , with at least one of the inequalities being strict.

Introducing a forward market increases aggregate production 12% in markets with three competitors ( $q(M3F) > q(M3)$ , p-value < 0.001). This confirms earlier findings such as in Brandts et al. (2008). Adding one more competitor in markets with three competitors increases aggregate production by 4%, and this effect is barely significant (p-value=0.105). We find that introducing a forward market increases aggregate production by 7% more than increasing competition by adding one more competitor, and this difference is strongly significant ( $q(M3F) > q(M4)$ , p-value=0.002).

**Results testing Hypothesis q.2: In markets with 2 competitors, both introducing a forward market and adding one more competitor increases production, and the strength of the effects are of the same order,  $q(M2F) > q(M2)$ ,  $q(M3) > q(M2)$ , and  $q(M2F) = q(M3)$ .**

We find support for Hypothesis q.2:

- $q(M2F) \leq q(M2)$  is REJECTED in favor of  $q(M2F) > q(M2)$ , p-value= 0.003.
- $q(M3) \leq q(M2)$  is REJECTED in favor of  $q(M3) > q(M2)$  , p-value= 0.006.
- $q(M2F) = q(M3)$  is NOT rejected in favor of  $q(M2F) \neq q(M3)$  , p-value= 0.374.
- $q(M2F) = q(M3) = q(M2)$  is REJECTED in favor of  $q(M2F) \geq q(M3) \geq q(M2)$ , with at least one of the inequalities being strict, p-value= 0.0000.

In line with the theoretical predictions, introducing a forward market increases aggregate production by 18% in markets with two competitors. This increase is strongly significant ( $q(M2F) > q(M2)$ , p-value= 0.003). Adding one more competitor in markets with two competitors increases aggregate production by 12%. This increase is significant ( $q(M3) > q(M2)$ , p-value= 0.006). Introducing a forward market increases aggregate production by 5% more than adding one more competitor, but this effect is not significant ( $q(M2F)=q(M2)$ , p-value= 0.344). A Jonckheere test rejects q.1 in favor of  $q(M2F) \geq q(M3) \geq q(M2)$ , p-value= 0.0000), with at least one of the inequalities being strict.

**Results testing Hypothesis q.3: Adding two more competitors does not increase production more than adding a forward market,  $q(M4) \leq q(M2F)$ .**

We find no support for Hypothesis q.3:

- $q(M4) \leq q(M2F)$  is NOT rejected in favor of  $q(M4) > q(M2F)$ , p-value= 0.521.

Doubling the number of competitors does not increase production significantly more than introducing a forward market. This is in contrast with the theoretical predictions. Our data indicate the opposite ordering instead;  $q(M2F)$  is 4% higher than  $q(M4)$ . This is surprising as LeCoq and Orzen (2006) found that the production of two competitors *with* forward market is strictly lower than that of four competitors *without* a forward market. The main difference is that in our treatments producers have steeply increasing production costs, while in their treatments producers have zero costs. Our result thus suggests that production costs make it harder for producers to collude. Indeed, in our treatments with two producers with production costs, M2 and M2F, subjects produced more than the Nash-equilibrium, while in our treatment with two producers and zero costs, M2zc and M2Fzc, subjects produced fewer units than the Nash-equilibrium. See the Appendix for a detailed discussion of our results for M2zc and M2Fzc.

**3.2. Efficiency**

As mentioned above, we define efficiency, following Brandts et al. (2008), as the joint consumer and producer surplus realized in the experiment divided by the maximum joint consumer and producer surplus (the Walrasian level of joint surplus). For the markets with a forward market, these measures are based on the outcomes in the forward and spot market together.

Figure 2 shows the evolution of efficiency per period, averaged over groups. Efficiency quickly converges and after period 8 its level is equal or higher than 90% for all treatments except M2. The highest efficiency levels in the last twelve periods are realized by treatments with forward markets, M2F and M3F.<sup>37</sup>

**Figure 2: Efficiency percentages**

a) M2, M2F, M3	b) M3, M3F, M4
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<sup>37</sup> See the Appendix for graphs of efficiency levels per period for the individual treatment together with the Nash-equilibrium prediction.

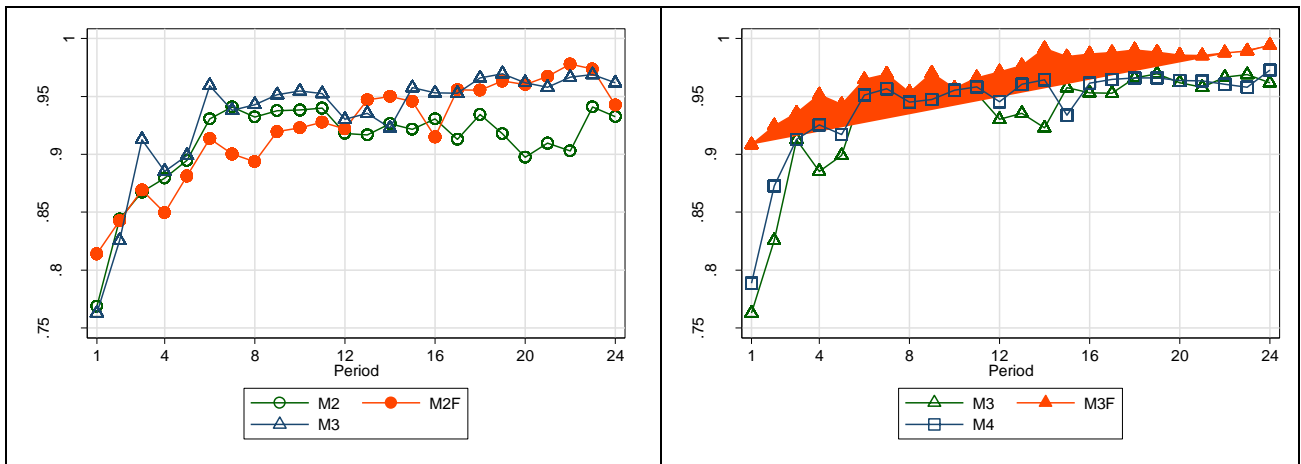


Table 7 shows the observed average efficiency level in the last 12 rounds, with the standard error in parenthesis. The row below gives the level of the observed average efficiency level relative to the Nash-equilibrium prediction in percentages. The efficiency levels are close to the Nash-equilibrium prediction; efficiency is significantly lower in M2 (p-value <0.068) and higher in M2F (p-value =0.083 in the low and 0.790 in the high Nash-equilibrium). This is mostly in line with earlier findings such as those in Brandts et al. (2008).

**Table 7: Efficiency averages in the last 12 rounds**

	M2	M2F	M3	M3F	M4
Average efficiency as % of Walras	92.0 (1.71)	95.5 (1.73)	95.6 (0.77)	98.7 (0.32)	96.1 (0.57)
% of NE prediction	97.2%	97.5%/ 100.7% <sup>38</sup>	98.3%	100.5%	98.6%
	N=11	N=11	N=11	N=11	N=11
% of NE prediction - earlier studies <sup>39</sup>	92.5%, LeCoq and Orzen (2006)	93,6%, LeCoq and Orzen (2006)	94.2%, Brandts et al. (2008)	96.7%, Brandts et al. (2008)	95.4%, Brandts et al. (2008) 109.3%, LeCoq and Orzen (2006)

Table 7 presents the results of the F-tests and Jonckheere test.<sup>40</sup> Aggregate production in the market is the most important determinant of efficiency, as production inefficiency has a minor influence only. The results of the tests of hypotheses regarding efficiency thus closely follow those regarding aggregate production.

**Table 8: Effects of one more competitor and forward market on efficiency, H $\Omega$ .1, H $\Omega$ .2 and H $\Omega$ .3**

	OLS regression, with correction for clustering on	Jonckheere test
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<sup>38</sup> The first number gives the percentage of efficiency relative to the high production Nash-equilibrium, the second number relative to the low production Nash-equilibrium.

<sup>39</sup> Using a Wilcoxon signed-rank test to compare with the results reported by Brandts et al. (2008) shows that in our results efficiency is significantly higher (p-values=0.003 for M3, M3F and M4). Compared with LeCoq and Orzen (2006), efficiency is significantly higher in condition M2F (p-value= 0.062 for the low Nash-equilibrium and p-value= 0.050 for the high Nash-equilibrium), significantly lower in condition M4 (p-value=0.003) and not significantly different in M2 (p-value= 0.131).

<sup>40</sup> The robustness tests, one-sided Wilcoxon rank-sum tests, confirmed our results at the same significance levels.



	group level, followed by a one-sided F-test on equality of the coefficients			
H $\Omega$ .1 - Markets with 3 producers	$\Omega(M3F) > \Omega(M3)^{***}$ (p<0.001)	$\Omega(M4) > \Omega(M3)$ (p=0.293)	$\Omega(M3F) > \Omega(M4)^{***}$ (p<0.001)	$\Omega(M3F) \geq \Omega(M4) \geq \Omega(M3)$ , with at least one of the inequalities being strict p-value < 0.001.
Number of observations	N= 792	N= 924	N= 924	N= 1320

H $\Omega$ .2 - Markets with 2 producers	$\Omega(M2F) > \Omega(M2)^*$ (p=0.075)	$\Omega(M3) > \Omega(M2)^{**}$ (p=0.026)	$\Omega(M2F) = \Omega(M3)$ (p= 0.927)	$\Omega(M2F) \geq \Omega(M3) \geq \Omega(M2)$ , with at least one of the inequalities being strict*** p-value < 0.001.
Number of observations	N= 528	N= 660	N= 660	N= 924

H $\Omega$ .3	$\Omega(M4) > \Omega(M2F)$ (p=0.351)
Number of observations	N= 792

**Results testing Hypothesis  $\Omega$ .1: In markets with 3 competitors, introducing a forward market increases efficiency, and the effect is stronger than adding one more competitor,  $\Omega(M3F) > \Omega(M4)$ , and  $\Omega(M3F) > \Omega(M3)$ .**

We find partial support for Hypothesis  $\Omega$ .1:

- $\Omega(M3F) \leq \Omega(M3)$  is REJECTED in favor of  $\Omega(M3F) > \Omega(M3)$ , p-value<0.001.
- $\Omega(M4) \leq \Omega(M3)$  is NOT rejected in favor of  $\Omega(M4) > \Omega(M3)$ , p-value=0.293.
- $\Omega(M3F) \leq \Omega(M4)$  is REJECTED in favor of  $\Omega(M3F) > \Omega(M4)$ , p-value<0.001.
- $\Omega(M3F) = \Omega(M4) = \Omega(M3)$  is REJECTED in favor of  $\Omega(M3F) \geq \Omega(M4) \geq \Omega(M3)$ , with at least one of the inequalities being strict, p-value<0.001.

Introducing a forward market in a market with three producers increases efficiency by 3.1% and this is strongly significant  $\Omega(M3F) > \Omega(M3)$ , p-value < 0.001). Adding one more competitor increases efficiency with a mere 0.5%, and this is not significant (NOT  $\Omega(M4) > \Omega(M3)$ , p-value = 0.293). The increase in efficiency from introducing a forward market is larger than that from adding one more competitor, and that effect is strongly significant ( $\Omega(M3F) > \Omega(M4)$ , p-value < 0.001).

**Results testing Hypothesis  $\Omega$ .2: In markets with 2 competitors, both introducing a forward market and adding one more competitor increases efficiency, and the strength of the effects are of the same order,  $\Omega(M2F) > \Omega(M2)$ ,  $\Omega(M3) > \Omega(M2)$ ,  $\Omega(M2F) > \Omega(M3)$ .**

We find support for Hypothesis  $\Omega$ .2:

- $\Omega(M2F) \leq \Omega(M2)$  is REJECTED in favor of  $\Omega(M2F) > \Omega(M2)$ , p-value=0.075.
- $\Omega(M3) \leq \Omega(M2)$  is REJECTED in favor of  $\Omega(M3) > \Omega(M2)$ , p-value=0.026.
- $\Omega(M2F) = \Omega(M3)$  is NOT rejected in favor of  $\Omega(M2F) \neq \Omega(M3)$ , p-value=0.927.

- $\Omega(M2F) = \Omega(M3) = \Omega(M2)$  is REJECTED in favor of  $\Omega(M2F) \geq \Omega(M3) \geq \Omega(M2)$ , with at least one of the inequalities being strict,  $p\text{-value} < 0.001$ .

Introducing a forward market increases efficiency by 3.5% and this is significant ( $\Omega(M2F) > \Omega(M3)$ ,  $p\text{-value} = 0.075$ ). Adding one more competitor increases efficiency with 1.1% and this is also significant ( $\Omega(M3) > \Omega(M2)$ ,  $p\text{-value} = 0.026$ ). The increase in efficiency due to the introduction of a forward market is not significantly larger than that due to adding one more competitor (NOT  $\Omega(M3F) \neq \Omega(M4)$ ,  $p\text{-value} = 0.927$ ).

**Results testing Hypothesis  $\Omega.3$ : Adding two more competitors does not increase efficiency more than introducing a forward market,  $\Omega(M2F) \leq \Omega(M4)$ .**

We find no support for Hypothesis  $\Omega.3$ :

- $\Omega(M4) \leq \Omega(M2F)$  is NOT rejected in favor of  $\Omega(M4) > \Omega(M2F)$ ,  $p\text{-value} = 0.351$ .

The effect of introducing a forward market with two competitors does not increase efficiency significantly less than doubling the number of competitors.

**3.3. Production Efficiency**

As mentioned above, we define production efficiency, following Brandts et al. (2008), as the actual producer surplus divided by the producer surplus had production taken place in the most efficient manner.<sup>41</sup> Figure 3 shows the evolution of efficiency per period, averaged over groups. Efficiency quickly converges and after period 8 its level is mostly equal or higher than 90% for all treatments.

The treatments with two traders are represented by circles, with three traders by triangles, and with four traders by squares. The treatments without forward markets are represented by open rounds, triangles or squares, the treatments with forward markets by filled rounds or triangles. M3 is clearly lower than M2, and M2F is most of the time in the middle. M4 is clearly lower than M3 and M3F, while there is no visible difference between M3 and M3F.

**Figure 3: Production Efficiency**

a) M2, M2F, M3	b) M3, M3F, M4
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<sup>41</sup> Given the quadratic marginal cost function this implies an as even as possible division of units over the producers.

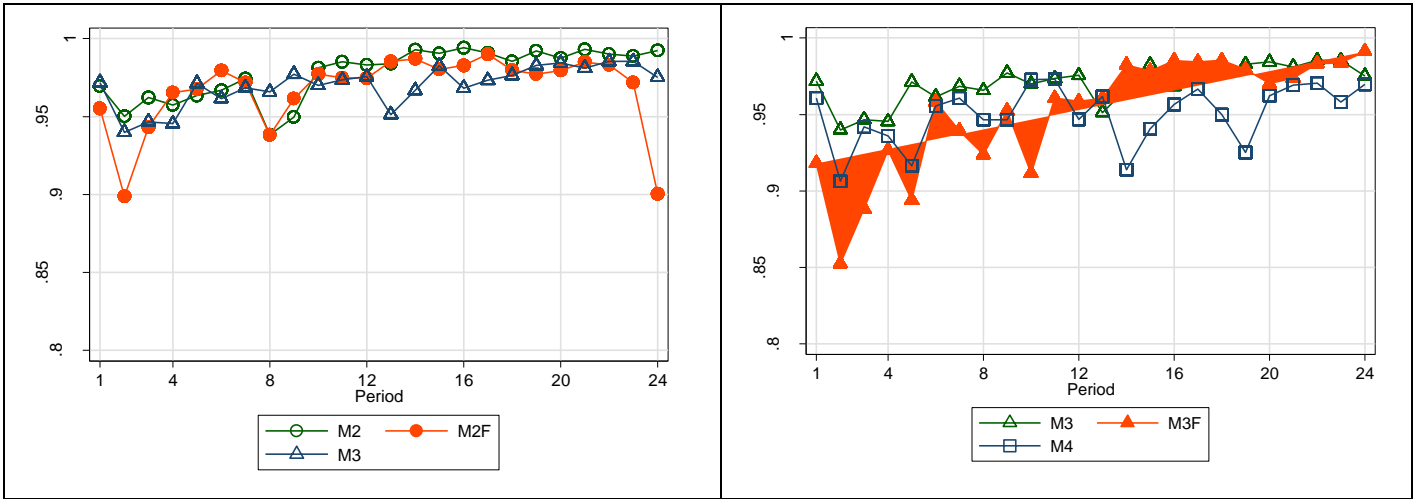


Table 9 shows the overall average of production efficiency in the last 12 rounds, with the standard error in parenthesis. The row below gives the size of the observed aggregated quantity relative to the Nash-equilibrium prediction in percentages.

**Table 9: Production efficiency averages in the last 12 rounds**

	M2	M2F	M3	M3F	M4
Average Production Efficiency	99.0 (0.35)	97.5 (0.81)	97.6 (0.59)	98.0 (0.69)	95.4 (1.63)
Number of observations	N=11	N=11	N=11	N=11	N=11

**Table 10: Effects of one more competitor and forward market on productive efficiency,  $H\Phi.1$  and  $H\Phi.2$**

	OLS regression, with correction for clustering on group level, followed by a one-sided F test	
$H\Phi.1$ – Markets with 3 producers	$\Phi(M4) < \Phi(M3)^*$ ( $p=0.093$ )	$\Phi(M3F) < \Phi(M3)$ ( $p=0.666$ )
Number of observations	N= 1001	N= 858
$H\Phi.2$ – Markets with 2 producers	$\Phi(M3) < \Phi(M2)^{**}$ ( $p=0.019$ )	$\Phi(M2F) < \Phi(M2)^{**}$ ( $p=0.046$ )
Number of observations	N= 715	N= 572

Table 10 presents the test for our hypothesis using F-tests based on an OLS regression and Jonckheere tests.<sup>42</sup>

**Results testing Hypothesis  $\Phi.1$ : In markets with 3 competitors, introducing a forward market does not decrease productive efficiency, while adding one more competitor does,  $\Phi(M4) < \Phi(M3)$  and  $\Phi(M3F) \geq \Phi(M3)$ .**

We find support for Hypothesis  $\Phi.1$ :

- $\Phi(M4) \geq \Phi(M3)$  is REJECTED in favor of  $\Phi(M4) < \Phi(M3)$  ,  $p$ -value=0.093.

<sup>42</sup> Robustness tests confirm our results, but show a weaker significance ( $p$ -value=0.100) for  $\Phi(M4) < \Phi(M3)$ .

- $\Phi(M3F) \geq \Phi(M3)$  is NOT rejected in favor of  $\Phi(M3F) < \Phi(M3)$  , p-value=0.666.

Adding one more competitor to M3 decreases the production efficiency by 2.4%, and this decrease is significant ( $\Phi(M4) < \Phi(M3)$  , p-value=0.093). Introducing a forward market does not lower production efficiency; the data rather suggest the opposite as efficiency is higher in the market with a forward market than in the market without one (though not significantly so).

**Results testing Hypothesis  $\Phi.2$ : In markets with 2 competitors, introducing a forward market and adding one more competitor decrease productive efficiency,  $\Phi(2F) < \Phi(M2)$ , and  $\Phi(3F) < \Phi(M3)$**

We find support for Hypothesis  $\Phi.2$ :

- $\Phi(M3) \geq \Phi(M2)$  is REJECTED in favor of  $\Phi(M3) < \Phi(M2)$  , p-value=0.019.
- $\Phi(M2F) \geq \Phi(M2)$  is REJECTED in favor of  $\Phi(M2F) < \Phi(M2)$  , p-value=0.046.

Adding one more competitor to M2 decreases production efficiency by 1.4%.<sup>43</sup> Introducing a forward market to a market decreases production efficiency by 1.5%. Both decreases are significant.

**3.4 Rationality in forward markets**

Using the assumption of rational behavior, Allaz and Vila (1993) argue that the forward price will be equal to the spot price, which implies that traders make zero profits. We indeed see this in our data for the treatments with a forward market: M2F and M3F. We estimated the relative markup of the spot market over the forward market price, defined by the difference between the two,

divided by the average price:  $P_{S-F} = \frac{P_S - P_F}{\frac{1}{2}(P_S + P_F)}$ . The average of  $P_{S-F}$  over the last 12 rounds is

0.001, which is not significantly larger than zero (p<0.97). This indicates that traders are making an insignificantly small profit. The total number of units producers offer on the forward market thus accurately predicts the total number of units they sell on the spot market, which indicates rational behavior.

**3.5 Summary of results and comparison to earlier experiments**

Table 11 summarizes our theoretical and experimental results for aggregate production, together with the key results of earlier experiments. We do not summarize the data on efficiency and productive inefficiency because the data on efficiency closely follow the patterns of the data on aggregate production, while the effect of productive inefficiency is small and inconsequential (see section 3.3).

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<sup>43</sup> Running, in addition, a Jonckheere test rejects  $\Phi(M4) \leq \Phi(M3) \leq \Phi(M2)$  in favor of  $\Phi(M4) \leq \Phi(M3) \leq \Phi(M2)$  , with at least one of the inequalities being strict, p-value=0.0000.

**Table 11: Comparison of our results with those of earlier studies**

		Theoretical predictions in our study	Results of earlier studies	Our study
Market with 2 competitors	One more competitor	+ 7.5%	-	+ 12.1% **
	FM	<ul style="list-style-type: none"> <li>• Same (low Nash-equilibrium)</li> <li>• + 10% (high Nash-equilibrium)</li> </ul>	+ 20.9% *** (LeCoq&Orzen, 2006)	+ 17.8% ***
	<b>Largest increase by</b>	<ul style="list-style-type: none"> <li>• <b>One more Competitor:</b> 7.5% higher than FM (low Nash-equilibrium)</li> <li>• <b>Forward Market:</b> 2.3% higher than OMC (high Nash-equilibrium)</li> </ul>	-	<b>Forward Market:</b> 4.7% higher than OMC (not significant)

Market with 3 competitors	One more competitor	+ 2.3%	+ 19.6% *** (Brandts et al., 2008 )	+ 4.4% (not significant)
	FM	+ 4.7%	+ 9.5% ** (Brandts et al., 2008 )	+ 12.0% ***
	<b>Largest increase by</b>	<b>Forward Market:</b> 2.3% higher than One more competitor	<b>One more Competitor:</b> 9.2% higher than FM** (Brandts et al., 2008)	<b>Forward Market:</b> 7.3% higher than OMC***

◻ : Results contrast with earlier results

◼ : Results contradict earlier results

Our results show that in markets with three competitors, in line with our theoretical prediction and with earlier experimental results (Brandts et al., 2008), introducing a forward market significantly increases aggregate production. Introducing a forward market increases aggregate production significantly more than adding one more competitor, which is in line with our theoretical prediction, but contradicts the findings of Brandts et al. (2008) (the contradictory findings are indicated by the shaded background in Table 11). In line with our theoretical prediction, adding one more competitor increases aggregate production. The increase is, however, not significant, which is in contrast with the findings of Brandts et al. (2008). The lack of significance is likely caused by the relatively small number of observations.

In markets with two competitors, in line with earlier experimental results (LeCoq and Orzen, 2006), introducing a forward market significantly increases aggregate production. Our data suggest that this increase is larger than that of adding one more competitor: The difference is not significant

but has a marginal significance in our robustness test. The lack of significance is also likely caused by the relatively small number of observations.

Our present study contributes to an understanding of the effects of forward markets and competition in electricity market settings. The following questions for future work suggest themselves: Is the effect of a forward market the same in an experiment with an indefinite time horizon? Demand uncertainty also introduces an insurance motive on the side of producers. Does the strategic effect of forward markets stay the same under demand uncertainty?

#### **4. Conclusion**

We have tried to better understand the comparative advantages of structural remedies and behavioral remedies of deregulation in electricity markets. We investigate theoretically and experimentally the effects of the introduction of a forward market on competition in electricity markets. We compared this scenario with the best alternative, reducing concentration by adding one more competitor by divestiture. Our work contributes to the literature by introducing more realistic cost configurations of steeply increasing marginal costs, teasing apart competition effect and asset effect, and studying numbers of competitors that better reflect the market concentration in the European states.

Our experimental results suggest that the behavioral remedy of introducing a forward market in concentrated markets with two or three competitors is an effective remedy for increasing the aggregate supply. This is in line with the empirical studies of Wolak (2001) and Van Eijkel and Moraga-Gonzalez (2010), who found empirical evidence that forward trading increased the aggregate supply in the Australian power market and in the Dutch gas market, respectively. Our experimental results also suggest that the effect of the behavioral remedy of introducing a forward market is larger than that of the structural remedy of adding one more competitor by divestiture. This is a policy relevant finding: competition authorities should, in line with EU law, focus on the behavioral remedy of introducing a forward market than on the structural remedy of lowering market concentration by divestiture.

At present, the EU has no single policy for the design of forward markets for electricity. Such a policy might improve the effectiveness of forward markets in the EU, as design is an important factor for the thickness of forward markets in EU countries (European Commission, 2007a, p.127). In Spain, for example, forward trading is de facto forbidden by design (European Commission, 2007a, p.127). In Greece forward trading has been made virtually impossible by design, as it has made trading in the pool mandatory (European Commission, 2007b, p.50). In contrast, in France the PowerNext exchange market allows for the trading of forward and future contracts of months,

quarters, and years ahead. Our study indicates that the design, or evolution, of such public forward exchanges as in France (and many other developed markets) should indeed be encouraged. Moreover, the public observability of forward positions is essential for the competition-increasing effect of Allaz and Vila (1993) to arise (Hughes and Kao, 1997; Van Eijkel and Moraga-Gonzalez, 2010). Observability of forward positions may not be optimal in markets with large volumes of over-the-counter trading. The EU could thus implement methods to increase the observability of the forward positions, for example by having the regulator publish aggregated and anonymized totals of forward positions.

Our results contradict the findings of Brandts et al. (2008), who found a stronger effect for the structural remedy of adding one more competitor than for the behavioral remedy of introducing a forward market. Their result stems most likely from the confound of competition effect and asset effect. In Brandts et al. (2008) adding one more competitor not only increases competition, but also increases the aggregate asset base, which reduces the aggregate cost and thus gives an extra incentive to increase production. This asset effect is likely influential, as producers have steeply increasing costs. The welfare effects Brandts et al. (2008) report are not conclusive, however, as they do not incorporate the costs of the increase in the asset base (the cost of building extra production plants). In our study we control for the asset effect by adding one more competitor by divestiture. As a result the effect of the structural remedy of adding one more competitor is weaker and is now dominated by the effect of the behavioral remedy of introducing a forward market.

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## 6. Appendices

### A1. Production costs

Table 12: Overview of aggregate cost of producing (rounded numbers)

Market with two producers (original market)			Market with three producers (after first divestment)					Market with four producers (after second divestment)						
Each Producer			Aggregate		Each Producer			Aggregate		Each Producer			Aggregate	
Units produced by each producer	Marginal Costs	Total Costs	Total Production	Total Costs	Units produced by each producer	Marginal Costs	Total Costs	Total Production	Total Costs	Units produced by each producer	Marginal Costs	Total Costs	Total Production	Total Costs
N	MC	TC	2*N	2*TC	N	MC	TC	3*N	3*TC	N	MC	TC	4* N	4*TC
0	0	0	0	0	0	0	0	0	0	0	0	0	0	0
1	1	1	2	2	1	2	2	3	6					
2	5	6	4	12						1	3	3	4	12
3	9	15	6	30	2	8	10	6	30					
4	16	31	8	62						2	12	15	8	60
5	24	55	10	110	3	18	28	9	84					
6	35	90	12	180	4	32	60	12	180	3	30	45	12	180
7	45	135	14	270	5	50	110	15	330					
8	60	195	16	390						4	55	100	16	400
9	80	275	18	550	6	70	180	18	540					
10	90	365	20	730						5	85	185	20	740
11	115	480	22	960	7	100	280	21	840					
12	130	610	24	1220	8	130	410	24	1230	6	120	305	24	1220
13	160	770	26	1540	9									
14	180	950	28	1900		160	570	27	1710	7	170	475	28	1900
15	210	1160	30	2320	10	200	770	30	2310					
16	230	1390	32	2780						8	220	695	32	2780
17	260	1650	34	3300	11	240	1010	33	3030					
18	300	1950	36	3900	12	290	1300	36	3900	9	280	975	36	3900
19	330	2280	38	4560	13	340	1640	39	4920					
20	360	2640	40	5280						10	345	1320	40	5280
21	410	3050	42	6100	14	390	2030	42	6090					
22	430	3480	44	6960						11	420	1740	44	6960
23	490	3970	46	7940	15	450	2480	45	7440					
24	520	4490	48	8980	16	510	2990	48	8970	12	500	2240	48	8960
25	560	5050	50	10100	17	580	3570	51	10710					
26	620	5670	52	11340						13	590	2830	52	11320
27	660	6330	54	12660	18	650	4220	54	12660					
28	710	7040	56	14080						14	690	3520	56	14080
29	760	7800	58	15600	19	720	4940	57	14820					
30	810	8610	60	17220	20	800	5740	60	17220	15	790	4310	60	17240
31	870	9480	62	18960	21	880	6620	63	19860					
32	920	10400	64	20800						16	890	5200	64	20800
33	1000	11400	66	22800	22	970	7590	66	22770					
34	1050	12450	68	24900						17	1010	6210	68	24840
35	1100	13550	70	27100	23	1060	8650	69	25950					
36	1150	14700	72	29400	24	1150	9800	72	29400	18	1140	7350	72	29400

37	1230	15930	74	31860	25	1250	11050	75	33150						
38	1320	17250	76	34500						19	1270	8620	76	34480	
39	1350	18600	78	37200	26	1350	12400	78	37200						
40	1450	20050	80	40100						20	1380	10000	80	40000	
41	1500	21550	82	43100	27	1450	13850	81	41550						
42	1600	23150	84	46300	28	1600	15450	84	46350	21	1550	11550	84	46200	
43	1650	24800	86	49600	29	1650	17100	87	51300						
44	1750	26550	88	53100						22	1700	13250	88	53000	
45	1800	28350	90	56700	30	1800	18900	90	56700						
46	1900	30250	92	60500						23	1900	15150	92	60600	
47	2000	32250	94	64500	31	1950	20850	93	62550						
48	2050	34300	96	68600	32	2050	22900	96	68700	24	2000	17150	96	68600	

## A2. Robustness tests

### A2.1 Alternate statistical tests

As robustness tests, we ran one-sided Wilcoxon rank-sum tests, as in LeCoq and Orzen (2006), for our hypotheses on quantity, efficiency and productive efficiency.

Table 13 shows the results of the robustness tests on quantity. Overall they confirm our findings in the main test with two exceptions. The relationship  $q(M4) > q(M3)$  is not significant anymore (p-value=0.154), but barely so. The relationship  $q(M2F) > q(M3)$  has a lower p-value and thus is significant (p-value= 0.086).

**Table 13: Test results quantity hypotheses**

	One-sided two-sample Wilcoxon rank-sum (Mann-Whitney) test		
Hq.1 - Markets with 3 producers	$q(M3F) > q(M3)^{***}$ ( $p < 0.001$ )	$q(M4) > q(M3)$ ( $p = 0.154$ )	$q(M3F) > q(M4)^{***}$ ( $p = 0.010$ )
	N= 22	N= 22	N= 22
Hq.2 - Markets with 2 producers	$q(M2F) > q(M2)^{**}$ ( $p = 0.01275$ )	$q(M3) > q(M2)^{**}$ ( $p = 0.012$ )	$q(M2F) > q(M3)^*$ ( $p = 0.070$ )
Number of observations	N= 22	N= 22	N= 22
Hq.3	$q(M4) > q(M2F)$ ( $p = 0.794$ )		
	N= 22		

Table 14 shows the results of the robustness tests on efficiency. Overall they confirm our findings in the main test; all relationships have the same levels of significance (0.1, 0.05, or 0.01) as in the main test.

**Table 14: Test results for H $\Omega$ .1, H $\Omega$ .2 and H $\Omega$ .3**

	One-sided two-sample Wilcoxon rank-sum (Mann-Whitney) test		
H $\Omega$ .1 - Markets with 3 producers	$\Omega(M3F) > \Omega(M3)^{***}$ (p= <b>0.002</b> )	$\Omega(M4) > \Omega(M3)$ (p= 0.311)	$\Omega(M3F) > \Omega(M4)^{***}$ (p< <b>0.001</b> )
Number of observations	N= 22	N= 22	N= 22

H $\Omega$ .2 - Markets with 2 producers	$\Omega(M2F) > \Omega(M2)^*$ (p=0.079)	$\Omega(M3) > \Omega(M2)^{**}$ (p= <b>0.039</b> )	$\Omega(M2F) > \Omega(M3)$ (p= 0.7251)
Number of observations	N= 22	N= 22	N= 22

H $\Omega$ .3	$\Omega(M4) > \Omega(M2F)$ (p=0.603)
Number of observations	N= 22

Table 15 shows the results of the robustness tests on production efficiency. Overall they confirm our findings in the main test with one exception: The relationship  $\Phi(M4) < \Phi(M3)^*$  has a slightly higher p-value and thus is no longer significant (p-value= 0.100), but barely so.

**Table 15: Test results for H $\Phi$ .1 and H $\Phi$ .2**

	One-sided two-sample Wilcoxon rank-sum (Mann-Whitney) test	
H $\Phi$ .1 – Markets with 3 producers	$\Phi(M4) < \Phi(M3)$ (p=0.100)	$\Phi(M3F) < \Phi(M3)$ (p= 0.859)
Number of observations	N= 22	N= 22

H $\Phi$ .2– Markets with 2 producers	$\Phi(M3) < \Phi(M2)^{**}$ (p= <b>0.041</b> )	$\Phi(M2F) < \Phi(M2)^*$ (p= <b>0.079</b> )
Number of observations	N= 22	N= 22

Notably, the robustness tests confirm the results we found in the main tests, and suggest that introducing a forward market may also have a stronger effect on competition than adding one more competitor in markets with two competitors.

## A2.2 Comparability data without costs

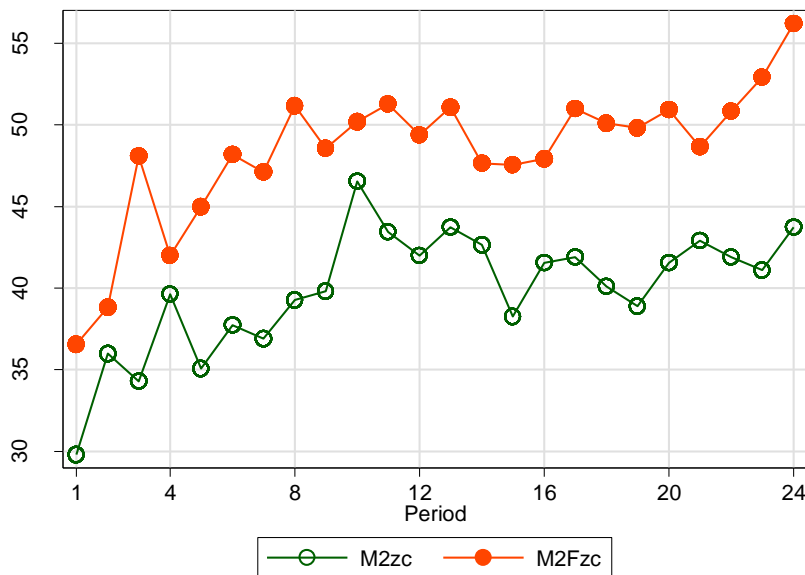
We ran treatments for markets with two producers without costs to allow comparisons with an earlier experiment on the effect of forward markets by LeCoq and Orzen (2006). Table 16 shows the theoretical predictions for these cases.

**Table 16: Theoretical predictions for zero-cost markets**

	NE M2-zc	NE M2F-zc	Walras-zc (n=2)	JPM-zc (n=2)
$q_{ii}^f$	–	16	–	–
$q_{ii}$	25	30	37	18/ 19 <sup>44</sup>
$q_t$	50	60	74	37
$p_t$	650	380	2	1001
<i>Prod. S.</i>	32500	22800	148	37037
<i>Cons. S.</i>	33075	47790	72927	17982
<i>Total S.</i>	65575	70590	73075	55019
<i>Eff. (%)</i>	89.74	96.60	100	75.29

Figure 4 shows the evolution of total (aggregate) quantities sold per period, averaged over groups. The treatments without forward markets are represented by open circles, the treatments with forward markets by filled circles. As in all other treatments, the aggregate production starts out rather low, 45, then quickly jump up in the direction of the Nash-equilibrium. Between rounds 10 and 12 behavior has stabilized.

**Figure 4: Average aggregate quantities sold per period**



<sup>44</sup> One generator produces 18 units, the other 19 units.

<sup>45</sup> We believe this might be a primer effect of the instructions, which presented examples with rather low numbers to facilitate understanding of the basic relationships.

### Averages by group

Table 17 shows that aggregate production tends to be significantly (p-values<0.093) smaller than the Nash-equilibrium, confirming the results of LeCoq and Orzen (2006).

**Table 17: Production averages and comparison**

Averages		
	M2zc	M2Fzc
Average production	41.6 (1.91)	50.3 9 (2.51)
% of NE prediction	79.9%	83.8%
Number of observations	N=11	N=11
% of NE prediction <i>LeCoq and Orzen (2006)</i>	93,2%,	93,8%,

Using a one-sided Wilcoxon rank-sum test we find that the increase in aggregate production due to a forward market is significant (p-value=0.014), confirming the results of LeCoq and Orzen (2006). A robustness tests confirms this finding.

**Table 18: Tests of effect of forward market**

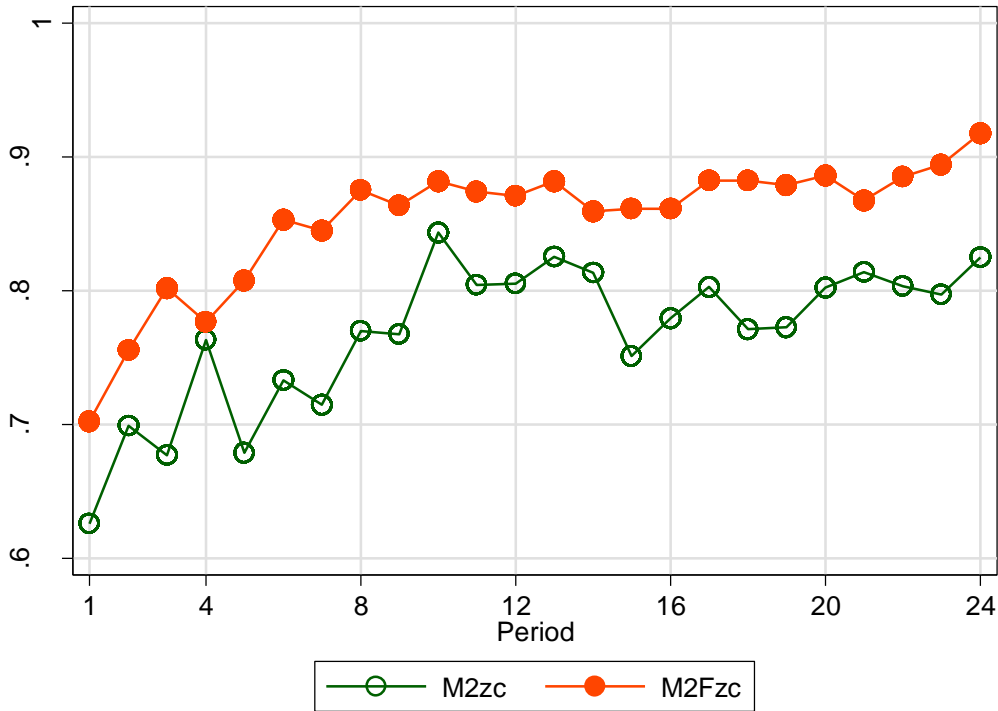
Main tests		
one-sided Wilcoxon rank-sum test		<b>M2Fzc &gt; M2zc**</b> <b>(p=0.014)</b>
		N=11

Robustness tests		
OLS regression with correction for clustering on group level, followed by one-sided F test on equality of the coefficients		<b>M2Fzc &gt; M2zc***</b> <b>(p&lt;0.010)</b>
		N= 572

Figure 5 shows the evolution of efficiency per period, averaged over groups. The treatments without forward markets are represented by open circles, the treatments with forward markets by filled circles. As producers have no production costs, production efficiency as defined in the main text is always 100%. Efficiency is thus determined by the aggregate production and the average efficiency in Figure 5 closely follows the aggregate average production (Figure 4).

**Figure 5: Average efficiency per period**



Efficiency is lower than the Nash-equilibrium prediction. A two-sided Wilcoxon one-sample signed-rank test indicates that these differences are significant (p-values<0.017).

**Table 19: Efficiency averages and comparison**

	M2zc	M2Fzc
Average efficiency as % of Walras	79.7 (2.10)	88.3 (2.37)
% of NE prediction	89.8%	90.7%
	N= 11	N= 11
one-sided Wilcoxon rank-sum test		<b>M2Fzc &gt; M2zc***</b> <b>(p&lt;0.010)</b>
		N= 16
OLS regression with correction for clustering on group level, followed by one-sided F test on equality of the coefficients		<b>M2Fzc &gt; M2zc**</b> <b>(p=0.011)</b>
	N= 572	N= 572

### A3. Sessions with experienced subjects

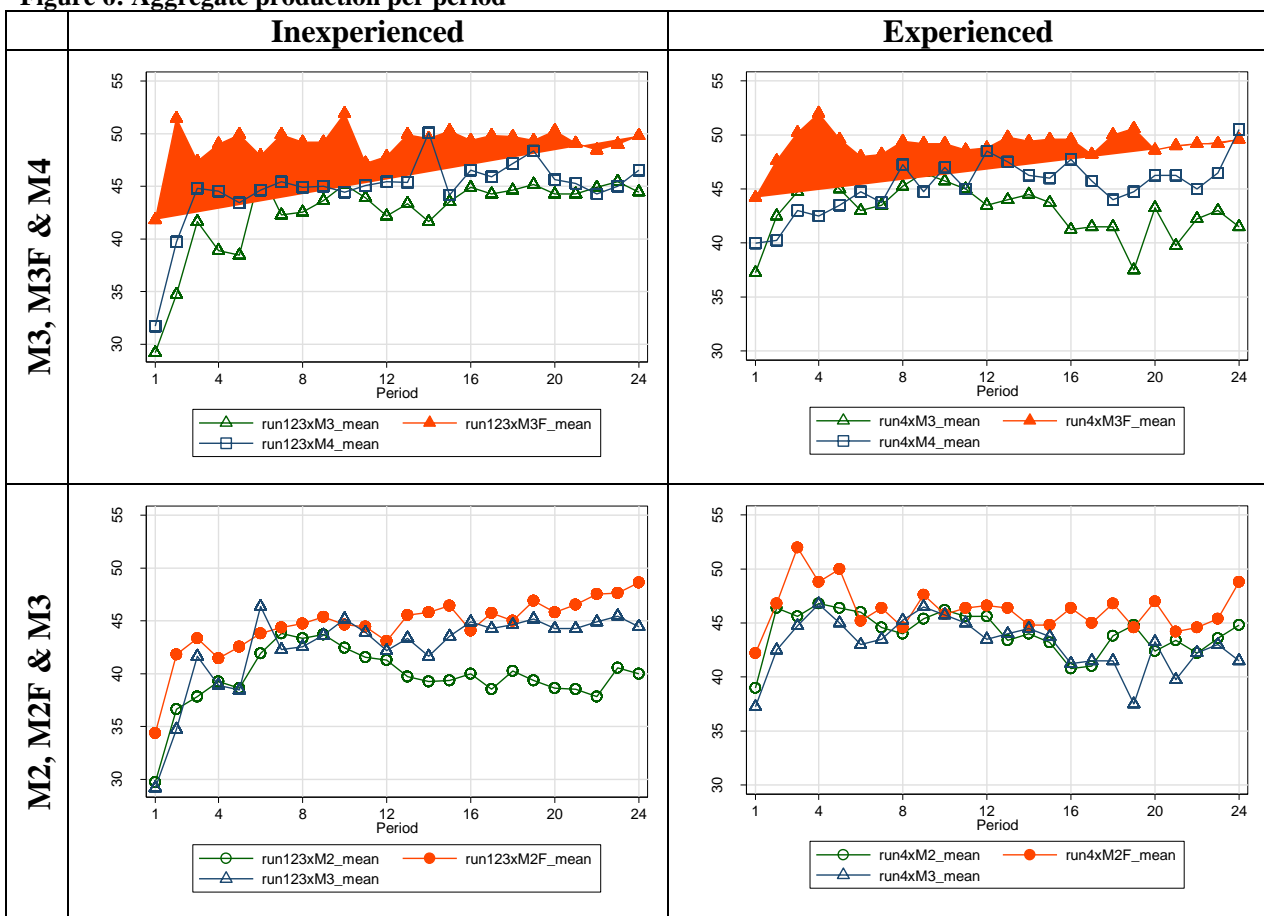
In October 2010 we ran sessions with subjects that had taken part in earlier sessions of the experiment. For each treatment we have 5 independent groups, except for M3F and M4, where we have 4 independent groups. For all treatments except M2 we assigned subjects to the exact same treatment they had participated in earlier. In the M2 treatments we had to rely on some subjects that had earlier participated in M3 or M4. Figure 6 shows the aggregate production for inexperienced subjects on the left and for experienced on the right. As can be seen, the basic pattern is the same.



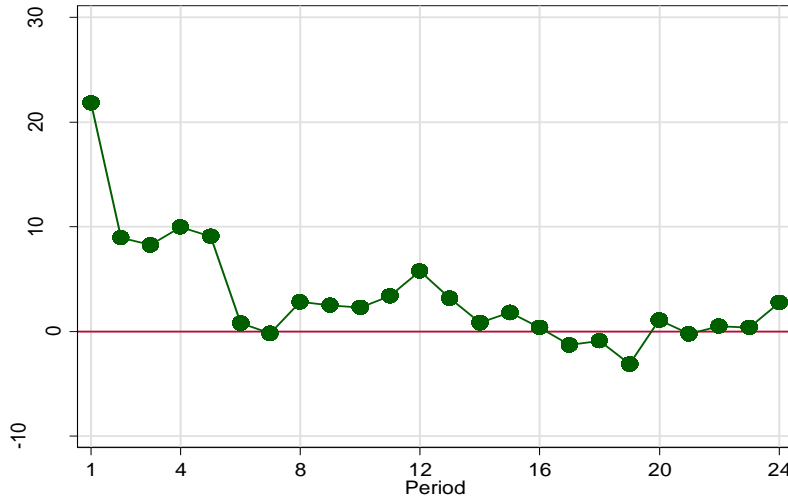
The eyeball test reveals that the largest differences are in the first periods – experienced subjects produce visibly more.

We noted above that subjects start out producing rather low (below the average over all rounds) and that we believe that this is a primer effect. In the instructions we gave examples with low numbers to facilitate understanding. We believe this triggered our subjects to produce low amounts in the first few rounds. Experienced subjects seem less susceptible to this instructional framing effect. Figure 7 shows how much more experienced subjects produced than inexperienced ones on average over all treatments. Experienced subjects produce 20% more in the first round. After that the difference levels off to 10%, and from round 6 on the difference hovers close to zero.

**Figure 6: Aggregate production per period**



**Figure 7: Percentage increase of aggregate production by experienced subjects**



To test if the aggregate production of experienced subjects is different from that of inexperienced subjects, we run a regression of the aggregate production on the treatment dummies interacted with the experience dummy for the last 12 periods:

$$AggSupply = \left\{ \begin{array}{l} InExp \cdot (\beta_{IE,M2} \cdot M2 + \beta_{IE,M2F} \cdot M2F + \beta_{IE,M3} \cdot M3 + \beta_{IE,M3F} \cdot M3F + \beta_{IE,M4} \cdot M4) + \\ Exp \cdot (\beta_{E,M2} \cdot M2 + \beta_{E,M2F} \cdot M2F + \beta_{E,M3} \cdot M3 + \beta_{E3F} \cdot M3F + \beta_{E,M4} \cdot M4) + \varepsilon \end{array} \right\}$$

$InExp$  ( $Exp$ ) is the dummy for the inexperienced (experienced) subjects, and  $M2$ ,  $M2F$ ,  $M3$ ,  $M3F$ , and  $M4$  are the dummies for the treatments. We then run 2-sided F-tests for the differences between  $\beta_{IE,i}$  (inexperienced subjects) and  $\beta_{E,i}$  (experienced subjects) for all treatments. Table 20 reports the results.

**Table 20: Aggregate production for experienced versus inexperienced subjects**

	M2	M2F	M3	M3F	M4
<b>Inexperienced</b>	39.34 (1.46)	46.31 (1.976)	44.24 (1.18)	49.55 (0.58)	46.20 (0.95)
<b>Experienced</b>	43.12 (1.53)	45.73 (2.40)	41.98 (1.57)	50.94 (0.16)	46.38 (0.93)
<b>Difference</b>	<b>+3.78</b>	-0.58	-2.26	<b>+1.39</b>	+0.18
Significance	<b>0.0781</b>	0.8533	0.2527	<b>0.025</b>	0.8937

Differences are significant for treatments M2 and M3F. In both these treatments experienced subjects produced more than inexperienced subjects. Another indication that experience leads to more production comes from regressing the aggregate production on the treatment dummies and a time variable to record how many months ago it is that the subject participated in the experiment. We conjecture that the more recently a subject has participated, the more influential his experience still is, and thus the more strongly might the effect of experience on the subject's decisions be. The

time variable is indeed negative and significant ( $p < 0.035$ ): subjects that participated 6 months ago produce on average 2.2 units more than subjects who participated 12 months ago.

Experience increasing aggregate production is in line with the experimental evidence for public provision games. This evidence indicates that private contributions to a public good fall with experience (Ledyard, 1995). The private contribution that producers could make in Cournot competition is to exercise restraint in selling units on the forward and spot markets. Such a restraint softens competition in the spot market which results in a higher overall profit. Experienced producers thus make lower “public contributions” by not restraining themselves.

#### A4. Predictions of the spot market price by our automated traders

M2F-zc: Total Production Stage A, Predicted Total Production and Resulting (Spot) Price

Total Production Stage A	Predicted (NE) Aggregate Production	Predicted (NE) price	Total Production Stage A	Predicted (NE) Aggregate Production	Predicted (NE) price	Total Production Stage A	Predicted (NE) Aggregate Production	Predicted (NE) price
0	49.4	667	33	71.4	73	66	93.4	0
1	50.0	649	34	72.0	55	67	94.0	0
2	50.7	631	35	72.7	37	68	94.7	0
3	51.4	613	36	73.4	19	69	95.4	0
4	52.0	595	37	74.0	1	70	96.0	0
5	52.7	577	38	74.7	0	71	96.7	0
6	53.4	559	39	75.4	0	72	97.4	0
7	54.0	541	40	76.0	0	73	98.0	0
8	54.7	523	41	76.7	0	74	98.7	0
9	55.4	505	42	77.4	0	75	99.4	0
10	56.0	487	43	78.0	0	76	100.0	0
11	56.7	469	44	78.7	0	77	100.7	0
12	57.4	451	45	79.4	0	78	101.4	0
13	58.0	433	46	80.0	0	79	102.0	0
14	58.7	415	47	80.7	0	80	102.7	0
15	59.4	397	48	81.4	0	81	103.4	0
16	60.0	379	49	82.0	0	82	104.0	0
17	60.7	361	50	82.7	0	83	104.7	0
18	61.4	343	51	83.4	0	84	105.4	0
19	62.0	325	52	84.0	0	85	106.0	0
20	62.7	307	53	84.7	0	86	106.7	0
21	63.4	289	54	85.4	0	87	107.4	0
22	64.0	271	55	86.0	0	88	108.0	0
23	64.7	253	56	86.7	0	89	108.7	0
24	65.4	235	57	87.4	0	90	109.4	0
25	66.0	217	58	88.0	0	91	110.0	0
26	66.7	199	59	88.7	0	92	110.7	0
27	67.4	181	60	89.4	0	93	111.4	0
28	68.0	163	61	90.0	0	94	112.0	0
29	68.7	145	62	90.7	0	95	112.7	0
30	69.4	127	63	91.4	0	96	113.4	0
31	70.0	109	64	92.0	0			
32	70.7	91	65	92.7	0			

**M2F: Total Production Stage A, Predicted Total Production and Resulting (Spot) Price**

Total Production Stage A	Predicted (NE) Aggregate Production	Predicted (NE) price	Total Production Stage A	Predicted (NE) Aggregate Production	Predicted (NE) price	Total Production Stage A	Predicted (NE) Aggregate Production	Predicted (NE) price
0	40.0	921	33	47.3	723	66	66.0	218
1	40.2	915	34	47.5	717	67	67.0	191
2	40.4	909	35	47.7	711	68	68.0	164
3	40.6	903	36	48.0	705	69	69.0	137
4	40.9	897	37	48.2	699	70	70.0	110
5	41.1	890	38	48.4	693	71	71.0	83
6	41.3	884	39	48.6	688	72	72.0	56
7	41.6	878	40	48.8	682	73	73.0	29
8	41.8	872	41	49.0	676	74	74.0	2
9	42.0	866	42	49.3	670	75	75.0	0
10	42.2	860	43	49.5	664	76	76.0	0
11	42.5	854	44	49.7	659	77	77.0	0
12	42.7	848	45	49.9	653	78	78.0	0
13	42.9	842	46	50.1	647	79	79.0	0
14	43.1	836	47	50.3	641	80	80.0	0
15	43.3	830	48	50.5	636	81	81.0	0
16	43.6	824	49	50.7	630	82	82.0	0
17	43.8	818	50	51.0	624	83	83.0	0
18	44.0	812	51	51.2	619	84	84.0	0
19	44.2	806	52	52.0	596	85	85.0	0
20	44.5	800	53	53.0	569	86	86.0	0
21	44.7	794	54	54.0	542	87	87.0	0
22	44.9	788	55	55.0	515	88	88.0	0
23	45.1	782	56	56.0	488	89	89.0	0
24	45.3	776	57	57.0	461	90	90.0	0
25	45.6	770	58	58.0	434	91	91.0	0
26	45.8	764	59	59.0	407	92	92.0	0
27	46.0	758	60	60.0	380	93	93.0	0
28	46.2	752	61	61.0	353	94	94.0	0
29	46.4	746	62	62.0	326	95	95.0	0
30	46.7	740	63	63.0	299	96	96.0	0
31	46.9	734	64	64.0	272			
32	47.1	728	65	65.0	245			

**M3F: Total Production Stage A, Predicted Total Production and Resulting (Spot) Price**

<b>Total Production Stage A</b>	<b>Predicted (NE) Aggregate Production</b>	<b>Predicted (NE) price</b>	<b>Total Production Stage A</b>	<b>Predicted (NE) Aggregate Production</b>	<b>Predicted (NE) price</b>	<b>Total Production Stage A</b>	<b>Predicted (NE) Aggregate Production</b>	<b>Predicted (NE) price</b>
0	43.2	833	33	48.4	693	66	66.0	218
1	43.4	829	34	48.6	688	67	67.0	191
2	43.5	824	35	48.7	684	68	68.0	164
3	43.7	820	36	48.9	680	69	69.0	137
4	43.9	816	37	49.0	676	70	70.0	110
5	44.0	811	38	49.2	672	71	71.0	83
6	44.2	807	39	49.3	668	72	72.0	56
7	44.3	803	40	49.5	663	73	73.0	29
8	44.5	799	41	49.7	659	74	74.0	2
9	44.7	794	42	49.8	655	75	75.0	0
10	44.8	790	43	50.0	651	76	76.0	0
11	45.0	786	44	50.1	647	77	77.0	0
12	45.1	781	45	50.3	643	78	78.0	0
13	45.3	777	46	50.4	639	79	79.0	0
14	45.5	773	47	50.6	635	80	80.0	0
15	45.6	769	48	50.7	630	81	81.0	0
16	45.8	764	49	50.9	626	82	82.0	0
17	45.9	760	50	51.0	622	83	83.0	0
18	46.1	756	51	51.2	618	84	84.0	0
19	46.2	752	52	52.0	596	85	85.0	0
20	46.4	747	53	53.0	569	86	86.0	0
21	46.6	743	54	54.0	542	87	87.0	0
22	46.7	739	55	55.0	515	88	88.0	0
23	46.9	735	56	56.0	488	89	89.0	0
24	47.0	730	57	57.0	461	90	90.0	0
25	47.2	726	58	58.0	434	91	91.0	0
26	47.3	722	59	59.0	407	92	92.0	0
27	47.5	718	60	60.0	380	93	93.0	0
28	47.6	713	61	61.0	353	94	94.0	0
29	47.8	709	62	62.0	326	95	95.0	0
30	48.0	705	63	63.0	299	96	96.0	0
31	48.1	701	64	64.0	272			
32	48.3	697	65	65.0	245			

**A5. Sheets given to the subjects  
(M2, M2zc, M3, M4)**

**Total Production and Resulting Price**

<b>Production</b>	<b>Price/Unit</b>	<b>Production</b>	<b>Price/Unit</b>	<b>Production</b>	<b>Price/Unit</b>
0	2000	33	1109	66	218
1	1973	34	1082	67	191
2	1946	35	1055	68	164
3	1919	36	1028	69	137
4	1892	37	1001	70	110
5	1865	38	974	71	83
6	1838	39	947	72	56
7	1811	40	920	73	29
8	1784	41	893	74	2
9	1757	42	866	75	0
10	1730	43	839	76	0
11	1703	44	812	77	0
12	1676	45	785	78	0
13	1649	46	758	79	0
14	1622	47	731	80	0
15	1595	48	704	81	0
16	1568	49	677	82	0
17	1541	50	650	83	0
18	1514	51	623	84	0
19	1487	52	596	85	0
20	1460	53	569	86	0
21	1433	54	542	87	0
22	1406	55	515	88	0
23	1379	56	488	89	0
24	1352	57	461	90	0
25	1325	58	434	91	0
26	1298	59	407	92	0
27	1271	60	380	93	0
28	1244	61	353	94	0
29	1217	62	326	95	0
30	1190	63	299	96	0
31	1163	64	272		
32	1136	65	245		

(M2F, M2Fzc, M3F)

**Aggregate Production and Resulting Price in STAGE B**

Aggregate number of Units in Stage A+ B	Resulting Price in STAGE B	Aggregate number of Units in Stage A+ B	Resulting Price in STAGE B	Aggregate number of Units in SPOT Market	Resulting Price in STAGE B
0	2000	33	1109	66	218
1	1973	34	1082	67	191
2	1946	35	1055	68	164
3	1919	36	1028	69	137
4	1892	37	1001	70	110
5	1865	38	974	71	83
6	1838	39	947	72	56
7	1811	40	920	73	29
8	1784	41	893	74	2
9	1757	42	866	75	0
10	1730	43	839	76	0
11	1703	44	812	77	0
12	1676	45	785	78	0
13	1649	46	758	79	0
14	1622	47	731	80	0
15	1595	48	704	81	0
16	1568	49	677	82	0
17	1541	50	650	83	0
18	1514	51	623	84	0
19	1487	52	596	85	0
20	1460	53	569	86	0
21	1433	54	542	87	0
22	1406	55	515	88	0
23	1379	56	488	89	0
24	1352	57	461	90	0
25	1325	58	434	91	0
26	1298	59	407	92	0
27	1271	60	380	93	0
28	1244	61	353	94	0
29	1217	62	326	95	0
30	1190	63	299	96	0
31	1163	64	272		
32	1136	65	245		



(M3F)

**Total Production STAGE A and Resulting Price in STAGE A**

<b>Total production STAGE A</b>	<b>Price STAGE A</b>	<b>Total production STAGE A</b>	<b>Price STAGE A</b>	<b>Total production STAGE A</b>	<b>Price STAGE A</b>
0	833	33	693	66	218
1	829	34	688	67	191
2	824	35	684	68	164
3	820	36	680	69	137
4	816	37	676	70	110
5	811	38	672	71	83
6	807	39	668	72	56
7	803	40	663	73	29
8	799	41	659	74	2
9	794	42	655	75	0
10	790	43	651	76	0
11	786	44	647	77	0
12	781	45	643	78	0
13	777	46	639	79	0
14	773	47	635	80	0
15	769	48	630	81	0
16	764	49	626	82	0
17	760	50	622	83	0
18	756	51	618	84	0
19	752	52	596	85	0
20	747	53	569	86	0
21	743	54	542	87	0
22	739	55	515	88	0
23	735	56	488	89	0
24	730	57	461	90	0
25	726	58	434	91	0
26	722	59	407	92	0
27	718	60	380	93	0
28	713	61	353	94	0
29	709	62	326	95	0
30	705	63	299	96	0
31	701	64	272		
32	697	65	245		

(M2F)

**Total Production STAGE A and Resulting Price in STAGE A**

<b>Total production STAGE A</b>	<b>Price/unit STAGE A</b>	<b>Total production STAGE A</b>	<b>Price/unit STAGE A</b>	<b>Total production STAGE A</b>	<b>Price/unit STAGE A</b>
0	921	33	723	66	218
1	915	34	717	67	191
2	909	35	711	68	164
3	903	36	705	69	137
4	897	37	699	70	110
5	890	38	693	71	83
6	884	39	688	72	56
7	878	40	682	73	29
8	872	41	676	74	2
9	866	42	670	75	0
10	860	43	664	76	0
11	854	44	659	77	0
12	848	45	653	78	0
13	842	46	647	79	0
14	836	47	641	80	0
15	830	48	636	81	0
16	824	49	630	82	0
17	818	50	624	83	0
18	812	51	619	84	0
19	806	52	596	85	0
20	800	53	569	86	0
21	794	54	542	87	0
22	788	55	515	88	0
23	782	56	488	89	0
24	776	57	461	90	0
25	770	58	434	91	0
26	764	59	407	92	0
27	758	60	380	93	0
28	752	61	353	94	0
29	746	62	326	95	0
30	740	63	299	96	0
31	734	64	272		
32	728	65	245		

(M2Fzc)

**Total Production STAGE A and Resulting Price in STAGE A**

<b>Total production STAGE A</b>	<b>Price/unit STAGE A</b>	<b>Total production STAGE A</b>	<b>Price/unit STAGE A</b>	<b>Total production STAGE A</b>	<b>Price/unit STAGE A</b>
0	667	33	73	66	0
1	649	34	55	67	0
2	631	35	37	68	0
3	613	36	19	69	0
4	595	37	1	70	0
5	577	38	0	71	0
6	559	39	0	72	0
7	541	40	0	73	0
8	523	41	0	74	0
9	505	42	0	75	0
10	487	43	0	76	0
11	469	44	0	77	0
12	451	45	0	78	0
13	433	46	0	79	0
14	415	47	0	80	0
15	397	48	0	81	0
16	379	49	0	82	0
17	361	50	0	83	0
18	343	51	0	84	0
19	325	52	0	85	0
20	307	53	0	86	0
21	289	54	0	87	0
22	271	55	0	88	0
23	253	56	0	89	0
24	235	57	0	90	0
25	217	58	0	91	0
26	199	59	0	92	0
27	181	60	0	93	0
28	163	61	0	94	0
29	145	62	0	95	0
30	127	63	0	96	0
31	109	64	0		
32	91	65	0		

(M2, M2F)

Units Produced	Marginal Costs	Total Costs	Production Costs		
			Units produced	Marginal Costs	Total Costs
1	1	1	25	560	5050
2	5	6	26	620	5670
3	9	15	27	660	6330
4	16	31	28	710	7040
5	24	55	29	760	7800
6	35	90	30	810	8610
7	45	135	31	870	9480
8	60	195	32	920	10400
9	80	275	33	1000	11400
10	90	365	34	1050	12450
11	115	480	35	1100	13550
12	130	610	36	1150	14700
13	160	770	37	1230	15930
14	180	950	38	1320	17250
15	210	1160	39	1350	18600
16	230	1390	40	1450	20050
17	260	1650	41	1500	21550
18	300	1950	42	1600	23150
19	330	2280	43	1650	24800
20	360	2640	44	1750	26550
21	410	3050	45	1800	28350
22	430	3480	46	1900	30250
23	490	3970	47	2000	32250
24	520	4490	48	2050	34300

(M3, M3F)

**Production Costs**

<b>Units Produced</b>	<b>Marginal Costs</b>	<b>Total Costs</b>
0	0	0
1	2	2
2	8	10
3	18	28
4	32	60
5	50	110
6	70	180
7	100	280
8	130	410
9	160	570
10	200	770
11	240	1010
12	290	1300
13	340	1640
14	390	2030
15	450	2480
16	510	2990
17	580	3570
18	650	4220
19	720	4940
20	800	5740
21	880	6620
22	970	7590
23	1060	8650
24	1150	9800
25	1250	11050
26	1350	12400
27	1450	13850
28	1600	15450
29	1650	17100
30	1800	18900
31	1950	20850
32	2050	22900

(M4)

**Production Costs**

<b>Units produced</b>	<b>Marginal Costs</b>	<b>Total Costs</b>
0	0	0
1	3	3
2	12	15
3	30	45
4	55	100
5	85	185
6	120	305
7	170	475
8	220	695
9	280	975
10	345	1320
11	420	1740
12	500	2240
13	590	2830
14	690	3520
15	790	4310
16	890	5200
17	1010	6210
18	1140	7350
19	1270	8620
20	1380	10000
21	1550	11550
22	1700	13250
23	1900	15150
24	2000	17150

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